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## Peterson, Chris L.

COMPARATIVE DEMOGRAPHY OF FOUR POPULATIONS OF THE HELLBENDER, CRYPTOBRANCHUS ALLEGANIENSIS, IN THE OZARKS

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# COMPARATIVE DEMOGRAPHY OF FOUR POPULATIONS OF THE HELLBENDER, CRYPTOBRANCHUS <br> ALLEGANIENSIS, IN THE OZARKS 

```
A Dissertation
Presented to
the Faculty of the Graduate School University of Missouri-Columbia
```

In Partial Fulfillment of the Requirements for the Degree Doctor of Philosophy

by<br>Chris Peterson<br>Dr. Dean E. Metter<br>Dissertation Supervisor<br>May 1985

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The undersigned, appointed by the Dean of the Graduate Faculty, have examined a dissertation entitled

Comparative demography of four populations of the hellbender, Cryptobranchus alleganiensis, in the Ozarks
presented by Chris Peterson
a candidate for the degree of Doctor of Philosophy
and hereby certify that in their opinion it is worthy of acceptance


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```
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Dr. Dean E. Metter Dissertation Supervisor

## ABSTRACT


#### Abstract

Four Ozark populations of the hellbender, Cryptobranchus alleganiensis, were censused to examine variation in life history among populations of two subspecies. The populations studied were the $\underline{C}$. $\underline{a}$. bishopi in the Spring River in Arkansas, and the Eleven Point River in Missouri, and the C. a. alleganiensis in the Gasconade and Big Piney Rivers in Missouri. Densities varied from 1.0 to 6.3 hellbenders / $100 \mathrm{~m}^{2}$ in different collection sites. Fecundity was positively related to increased length of females in all populations. $\underline{C}$. a. bishopi produced more ova at any given size than $\underline{C}$. a. alleganiensis females, but differences existed in average adult size among rivers. Stomach contents of females killed for fecundity data indicated that crayfish was the major food item in all populations, although hellbender eggs were eaten in large numbers during spawning by Gasconade females. The number of crayfish caught per unit effort was positively correlated with estimated production by hellhenders in the collection


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sites. Growth of hellbenders decreased linearly with increased
total length. Growth of adults was greatest in the Spring
River followed by hellbenders in the Gasconade, Big Piney,
and Eleven Point Rivers. The Eleven Point adults were generally
much smaller than adults in the other three rivers. The age-
specific fecundity of the Spring River and Gasconade River
females was greater than that of the Big Piney and Eleven
Point females. Assuming a replacement rate of 1.0, all popu-
lations exhibited a Type III survivorship curve. The adult
female survivorship in the Spring and Gasconade Rivers may
be slightly higher than that of Big Piney and Eleven Point
females. It is suspected that differences in age-specific
fecundity and possibly survivorship are related to growth.
Because hellbenders have been reported to have genetic uniformity
throughout their range, it is believed that differences in
growth in different populations were due mostly to environmental
factors that differed among rivers, perhaps food availability.
Demographic comparisons provide little support for the existence
of two subspecies in the Ozarks. Where there issuitable habitat,
hellbender secondary production is probably at least as great
as that of the larger fish species in the rivers.
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## INTRODUCTION

Cryptobranchus alleganiensis is a large, long-lived salamander. There are two described subspecies. The hellbender, ㄷ. a. alleganiensis, inhabits clear, cool streams from southern New York south to northern Georgia and Alabama and west through Tennessee and the Ohio River Valley to north-flowing Ozark streams of the Missouri River drainage (Dundee, 1971). C. a. bishopi, the Ozark hellbender, is known from the Spring River in Arkansas (Dowling, 1957) and from the North Fork of the White, Current, and Eleven Point Rivers in Missouri (Firschein, 1951), all south-flowing streams. Wortham and Nickerson (1971) speculated that the two races became geographically separated in the Tertiary. Grobman (1943) described the bishopi taxon, noting differences between the two races in size of spiracular openings, dorsal and labial spotting, and development of the lateral line system in the pectoral region. According to Dundee and Dundee (1965), any other morphological differences are slight. Wortham and Nickerson (1971) found that serum protein patterns of C. a. bishopi and C. a. alleganiensis were distinguishable with regard to their electrical migration through polyacrylamide gels. They suggested, as had Grobman (1943), that the two taxa should be treated as separate species. Jerrett and Mays (1973) reported significant differences in erythrocyte size and number, monocyte size, blood pH , and hemoglobin concentration between one population of C. a. bishopi and one population of C. a. alleganiensis. However, Merkle et al. (1977) stated that the bishopi subspecies status was not warranted
on a biochemical basis, but suggested that the taxon should probably be retained because of the above morphological differences. They found only two of 24 presumptive loci were polymorphic in 137 hellbenders from 12 populations, including three populations in the Ozarks. One unique allele occurred in North Fork and Spring River C. a. bishopi in frequencies of 0.28 and 0.35 , respectively. At a different locus the other unique allele was present in a Pennsylvania population. The mean proportion of loci heterozygous per individual for a particular population (H) was 0.0 in nine populations. In the other three populations, $H$ was 0.002 . Many other urodele populations have greater variability in genetic frequencies than the hellbender. For six Eurycea lucifuga populations in the Ozarks, H values varied from 0.0 to $0.106(\overline{\mathrm{H}}=0.039)$ based on 12 presumptive loci (Merkle, 1975). Shaffer (1983) examined 29 to 32 proteins in seven populations of Ambystoma rosaceum and two populations of $\underline{A}$. tigrinum. H levels varied from 0.101 to $0.176(\bar{H}=0.139)$ for $A$. rosacum localities. For the two A. tigrinum samples, $H$ was 0.09 and 0.171 . Even in Plethodon ouachitae with its limited range, $H$ values ranged from 0.015 to 0.084 $(\overline{\mathrm{H}}=0.055)$ among 10 localities based on 23 genetic loci (Duncan and Highton, 1979).

Ecological comparisons of C . a. bishopi and C. a. alleganiensis began with Dundee and Dundee (1965). They found there was an overlap in the breeding season of the two races, but that the Ozark hellbender generally oviposited later and at a smaller initial breeding size. Nickerson and Mays (1973a; 1973b) gathered considerable information about the natural history of $C$. $a$. bishopi, mostly from the North Fork River, including population densities, biomasses, food habits,

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activity patterns, sex ratios, breeding habits, and differences in
size of individuals among populations. Topping and Ingersol (1981)
reported that North Fork Ozark hellbenders have a greater length-
specific fecundity than C. a. alleganiensis in the Niangua River,
Missouri. Coatney (1982) compared the home range of hellbenders
in the North Fork with the home range reported for Pennsylvania
hellbenders by Hillis and Bellis (1971). Peterson et al. (1983)
reported differences between the age and growth of North Fork
and Niangua hellbenders from data obtained by Taber et al. (1975)
from the latter population. Recently the usefulness of the growth
model of Peterson et al. (1983) for Ozark hellbenders was confirmed
by resampling the North Fork population and comparing the observed
growth of recaptures with predicted growth (Peterson et al., 1985).
    Despite the similar genetic frequencies reported by Merkle et
al. (1977), some demographic traits (growth, fecundity) are different
between subspecies in Missouri. Also, prior sampling by the author
and by Nickerson and Mays (1973a) indicated that differences exist
in the adult size of hellbenders even in populations within a
subspecies. This alone suggests that differences exist in fecundity
among populations because fecundity scales with size (Topping and
Ingersol, 1981). It is unknown, however, whether differences in
demography are as profound among populations within a subspecies
as between subspecies. Further, the effects of environmental
factors on hellbender life history are unknown. Also, because of
the pattern observed by Organ (1961) of greater mortality of the
young in increasingly more aquatic species in the salamander genus
Desmognathus, the determination of survivorship in the strictly
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## LITERATURE REVIEW

## Wet Weight Biomass and Production

Only a few studies have dealt with the importance of aquatic urodeles in terms of standing crop, even though they may be important predators in some aquatic habitats. Burton (1977) found that 1950 to 2600 adult red-spotted newts, Notophthalmus viridescens, inhabited 1 ha of Mirror Lake, New Hampshire. Assuming an average weight of 3 g for an adult newt (cf. Gill, 1979), the standing crop was 5.8 to $7.8 \mathrm{~kg} /$ ha. Neish (1971) estimated a biomass of $37 \mathrm{~kg} / \mathrm{ha}$ for second-year larvae and neotenous adult Ambystoma gracile throughout a 10-ha lake in British Columbia. Gelbach and Kennedy (1978) presented evidence of dominance of a beaver pond by a single urodele, Siren intermedia. They estimated a biomass of 380 to $570 \mathrm{~kg} \mathrm{/} \mathrm{ha} \mathrm{and}$ believed that seven species of insectivorous fish in the pond had a standing crop totaling no more than $420 \mathrm{~kg} /$ ha. Their estimate of production by $\underline{s}$. intermedia was $290 \mathrm{~kg} /$ ha / yr.

Nickerson and Mays (1973b) estimated a biomass of about 450 $\mathrm{kg} /$ ha of hellbenders in one riffle in the North Fork River; only suitable diurnal habitat was considered in the estimate of area. If the area of the entire riffle had been considered, the estimate would have been about $105 \mathrm{~kg} /$ ha. Peterson et al. (1983) estimated an average hellbender standing crop of 157 and $154 \mathrm{~kg} / \mathrm{ha}$ for two collection sites in the North Fork based on shoreline to shoreline area. The latter authors speculated that even higher densities existed in the Niangua River. Topping (1981 and personal communication) estimated the production of Niangua hellbenders to be 61 to 714
$\mathrm{kg} / \mathrm{ha} / \mathrm{yr}$ for 10 different sites. Only the limited area where hellbenders could be caught during daylight was considered in these estimates. Other than the production reported for Siren intermedia, Topping's estimates are the only estimates known to the author of production by a truly aquatic salamander. These figures suggest that the hellbender's secondary production may be comparable to many stream £ishes.

Waters (1977) reviewed the literature on production of fish in streams and lakes. Estimates of production of the lotic benthic fishes Cottus (3 species) and Etheostoma (2 species) varied from 8 to $431 \mathrm{~kg} \mathrm{/} \mathrm{ha} \mathrm{/} \mathrm{yr} \mathrm{and} 23$ to $52 \mathrm{~kg} \mathrm{/} \mathrm{ha} \mathrm{/} \mathrm{yr}, \mathrm{respectively}$. yearly production of stream cyprinids ranged from 0.1 to $915 \mathrm{~kg} / \mathrm{ha}$ depending on the species and locality. Annual production estimates for brook trout, Salvelinus fontinalis, and rainbow trout, Salmo gairdneri, in different rivers varied from 16 to $300 \mathrm{~kg} \mathrm{/} \mathrm{ha} \mathrm{and}$ 28 to $132 \mathrm{~kg} / \mathrm{ha}$, respectively. The annual production of brook trout, mottled sculpin (Cottus bairdi), and blacknose dace (Rhinichthys atratulus), the three dominant species in a headwater stream in Virginia, was estimated at three different sites (Neves and Pardue, 1983). Production estimates were 5.4 to $19.3 \mathrm{~kg} /$ ha for trout, 5.7 to $10.8 \mathrm{~kg} \mathrm{/} \mathrm{ha} \mathrm{for} \mathrm{sculpin}$,and 2.7 to $3.8 \mathrm{~kg} \mathrm{/} \mathrm{ha} \mathrm{for} \mathrm{dace}$. Growth

In most amphibians there appears to be a period of rapid growth preceding attainment of sexual maturity, followed by a period during which growth tends to be intermittent and gradually slows (Porter, 1972). However, growth has been reported to decrease at a constant


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Growth of hellbender larvae can only roughly be estimated from reports. Eggs hatch in late fall, about six weeks after fertilization, at about 25 to 30 mm TL (Bishop, 1941; Smith, 1907). Larvae 35 to 60 mm in length have been reported to be three to nine months old (Grenell, 1939). One-year-old larvae vary in length between 60 and 90 mm TL (Taber et al., 1975; Grenell, 1939; Smith, 1907). Hellbenders in their second year may be 90 to 130 mm TL (Nickerson and Mays, 1973a; Bishop, 1941; Smith, 1907).

Bishop (1943) reported that metamorphosis occurs when hellbenders are between 100 and 130 mm in length and are about 18 months old. Grenell (1939) stated that gill loss occurs at the end of the second year when larvae have reached a length of at least 120 mm . Goin and Goin (1971) concluded that transformation occurs when hellbenders are 1.5 years old and their length is 125 mm . Nickerson and Mays (1973a) also concluded that 125 mm TL is the approximate size of Ozark hellbenders when they lose their gills.

Age-specific Fecundity and Survivorship
Because of their secretive habits, the analysis of caudate population dynamics is difficult, and many species remain unstudied (Porter, 1972). Difficulties in aging individuals is also a major problem in determining age-specific fecundity and survivorship. Methods of determining age of urodeles have included the examination of skeletal annuli (e. g., Peabody, 1961; Senning, 1940), counting testicular lobes (e. g., Organ, 1961; Humphrey, 1922), the detection of discrete size classes (e. g., Hairston, 1983; Tilley 1973), or mark-recapture studies (Peterson et al., 1983; Tilley, 1980; Taber et

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al., 1975). No skeletal growth rings are present in hellbenders
(Peabody, 1961), nor have testis lobes been described (cf. Ratcliff,
1965; Branch, 1959). Dundee and Dundee (1965) stated that at least
three size groups are present in postmetamorphic hellbenders. However,
Taber et al. (1975) were unable to estimate the age of hellbenders
in large samples by length-frequency methods. The growth models
developed by Taber et al. (1975) and Peterson et al. (1983) were used
to produce length-age relationships for hellbenders. However, some
bias exists in an age-length key if overlap between successive ages is
present (cf. Westrheim and Ricker, 1978) as is the case with old
salamanders.
Although the relationship between female body length and fecundity
varies among urodeles (Porter, 1972), in general, the size of the female
correlates directly with her reproductive capacity (Salthe, 1969).
Blanchard (1928) found no correlation between fecundity and size of
fonale Plethodon cinereus in Michigan. Yet Nagel (1977) described a
weak correlation between fecundity and body length for the same species
in Tennessee, and Lotter (1978) reported a strong correlation between
the two variables for P. cinereus in Connecticut. Stebbins (1954) found
a slight positive relationship between number of eggs produced and size
of female Ensatina up to 70 mm snout-vent length but a decline in eggs
produced by larger females. Harrison (1967) found a direct correlation
between clutch size and body size in Desmognathus aeneus. Tilley (1968)
also described a positive correlation between ova produced and female
snout-vent length in four species of Desmognathus but a negative
correlation in D. monticola. Although larger D. ochrophaeus females
produce larger clutches, fertility rates were uncorrelated with age
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survivorship curves for five species of Desmognathus based on the number of testis lobes. The more terrestrial the species, the greater was the early survivorship. The more aquatic species had a Type II survivorship curve indicating a constant rate of mortality for males one to four years old (Organ, 1961). Spight (1967) presented a life table for Desmognathus fuscus males in North Carolina from hatching through seven years of age. Again testicular lobes were used in aging. This method was also used by Danstedt (1975) to estimate age-specific survivorship in the same species. Bell (1977) estimated an annual survival rate of about $50 \%$ for adult smooth newts, Triturus vulgaris. Female survivorship was a little greater than that of males. Gill (1978) produced a Type II survivorship curve for Notophthalmus viridescens. Male survival was slightly higher than female survival. Tilley (1980), assuming a constant survivorship rate after sexual maturity, calculated a mean generation time of 5.6 years and a longevity of at least 16 years for Desmognathus ochrophaeus. Hairston (1983) calculated a mean generation time of 9.8 years for Plethodon jordani and suggested a life span of as much as 30 years. Thus even rather small urodeles may have a relatively long life span.

Dundee and Dundee (1965) argued that hellbenders live at least seven years, and Nickerson and Mays (1973a) suggested that hellbenders are long-lived. Andrias has lived $55^{+}$years in captivity (Flower, 1936), and Cryptobranchus has survived 29 years in captivity (Oliver, 1955). Richmond (1965) hypothesized that hellbender populations are dominated by long-lived adults. Taber et al. (1975) confirmed this speculation for C. a. alleganiensis in the Niangua River, Missouri. They indicated a life span of at least 30 years and suggested that

```
adult female survivorship is greater than that of mature males. The
difference in survivorship between sexes is supported by the one male
to four female sex ratio reported by Wiggs (1977) for Niangua
hellbenders }15\mathrm{ years old or older. Peterson et al. (1983) in a
similar study on C. a. bishopi in the North Fork River reached similar
conclusions to those of Taber et al. (1975). Neither study combined
fecundity rates with survivorship to estimate mortality between
oviposition and sexual maturity.
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## STUDY AREAS <br> Spring River

The Spring River has its origin at Mammoth Spring, the second largest spring in the Ozarks (Beckman and Hinchey, 1944). It flows for about 114 km before joining the Black River. The spring is just south of the Missouri-Arkansas border in Fulton County. Nearly all of the drainage area of the spring is in Missouri. The discharge is about 152 million $1 / \mathrm{hr}$. Mammoth Spring is inundated by a dam within the town of Mammoth Spring (Ark. Dept. of Pollution Control and Ecology, 1977a; Vineyard and Feder, 1974).

Collection site 1 in the Spring River is located below Dam No. 3 about 7 km downstream from Mammoth Spring at the downstream end of a large island. A trout-raising operation is located on the island and utilizes water from the river for its silos. After settling in lagoons, the water is discharged back into the river just upstream from the collection site. Waste materials in this water include fish metabolic wastes and uneaten fish food (Ark. Dept. of Pollution Control and Ecology, 1977a). The river is still highly influenced by the spring at this point. Water temperatures are cool and fairly uniform throughout the year. Dundee and Dundee (1965) reported that the river for several miles downstream of the spring maintains a temperature of $15 \pm 2{ }^{\circ} \mathrm{C}$ throughout the year.

Site 2 in the Spring River is located on either side of an island about 1 km downstream from the dam at the spring. The east bank is a residential area. The sewage treatment plant for Mammoth Spring consists of an aerated lagoon and settling pond. The discharge is

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chlorinated before entering an unnamed tributary upstream from the collection site (Ark. Dept. of Pollution Control and Ecology, 1977b). The Warm Fork also enters upstream from the collection site. This tributary has about 15 km of permanent flow, mostly in Missouri. Sewage effluent from Thayer, Missouri, enters this creek (Duchrow, 1977). In 1976, fecal coliform and nitrates from municipal discharge into the Spring River were parameters not meeting the Federal Water Pollution Control Act and/or the Water Quality Standards of Arkansas (Ark. Dept. of Pollution Control and Ecology, 1977b).
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## Eleven Point River

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The Eleven Point River has its headwaters in Howell County Missouri. If flows for about 200 km and enters the Spring River near its junction with the Black River. Sixteen major springs contribute substantially to the flow (Vineyard and Feder, 1974). Greer Spring more than doubles the flow of the river and is Missouri's second largest spring. Although large areas of oak-hardwood forests in its headwaters have been converted to pasture, much of its length is in Mark Twain National Forest. The area is sparsely populated.
Duchrow (1977) made a water quality study of the Eleven Point and its major tributaries. The basin was considered unpolluted and unaffected by point sources of pollution. Regression analyses made by Duchrow (1977) for the sampling stations showed a positive relationship between the increase in nutrients and invertebrate species diversity index values. Enriched tributaries normally had higher macroinvertebrate densities. In most cases the tributaries with high species diversity index values had watersheds in which forest had been converted to pasture. However, no sources of pollution or timber clearing were
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known to exist from Greer Spring to a sampling station downstream of
the hellbender collection sites (Duchrow, 1977). Many springs
contribute to the river's flow in this stretch (Vineyard and Feder,
1974).
    Site 1 in the Eleven Point River is located about 12 km downstream
from Turner's Mill in Oregon County. Site 2 is located about 2 km
upstream from Turner's Mill or about 7 km downstream from the Greer
Spring Branch. Dundee and Dundee (1965) reported that river tempera-
tures remain below 17 '0}\textrm{C}\mathrm{ for many miles of its.course even in midsummer
and do not drop below 12.5 '}\textrm{C}\mathrm{ in midwinter.
                    Gasconade River
    The Gasconade River has its headwaters in Wright County in south-
central Missouri, and it flows northeast for about 440 km to its
junction with the Missouri River. It is subject to frequent and severe
flooding (U.S.G.S. and Mo. Div. of Geol. Surv. and Water Res., 1967).
Much of the upper portion of the Gasconade Basin is used extensively
for crop farming; however, in the hilly sections dairy farming is
predominant. Although towns are located in the basin, a }1973\mathrm{ report
indicated no improvements were needed in waste treatment in the area
(Mo. Water Pollution Bd., 1973). Although there are many springs
flowing into the river, according to Dundee and Dundee (1965), the
Gasconade is a larger stream than the south-flowing hellbender rivers,
and thus spring waters are not as effective in modifying its tempera-
tures. They recorded summer temperatures of 25 '}\textrm{C}\mathrm{ , and they speculated
that winter temperatures were perhaps colder than in the south-flowing
streams.
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Site 1 , the only collection site in the Gasconade River, is a long stretch about 1 km upstream from the highway 17 bridge in Pulaski County near Waynesville. Shorelines are generally muddy, and smooth bedrock areas are heavily covered with algae in summer. Following heavy rains this stretch remains turbid with low visibility, sometimes for many days. The sewage treatment plant at Waynesville discharges into Roubidoux Creek which joins the Gasconade at the highway 17 bridge downstream from the collection site.
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Big Piney River
The Big Piney River is located in the Gasconade Basin and flows into the Gasconade in Pulaski County 42 km downstream from the Gasconade collection site. The river originates in Texas County and flows through Fort Leonard Wood in Pulaski County. The middle portion of the stream runs through National Forest. No water treatment problems were noted at Fort Leonard Wood downstream from the collection site, nor at Houston, the only town upstream that could much affect the collection sites (Mo. Water Pollution Bd., 1973). However, water quality reports have indicated violation of dissolved oxygen standards due to municipal discharge from the town of Cabool in the headwaters of the Big Piney River (Mo. Clean Water Comm. Plan. Sec., 1982; 1980), but probably this did not affect the collection sites. As is typical of Ozark streams, the Big Piney River is exceptionally clear; turbidities were less than 7 ppm except immediately following heavy rains when measured by Fleener et al. (1974).

Site 1 in the Big Piney is located about 2.5 km upstream from the highway 32 bridge between Licking and Success in Texas County. Site 2 is located about 0.5 km downstream from the bridge at a privately
operated campground.

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## MATERIALS AND METHODS



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Thermometers were checked approximately once a month from fall 1980 to fall 1981, although flooding occasionally prevented readings, particularly in the Gasconade River. Thermometers were also checked in the North Fork River at the highway $H$ bridge in Ozark County, and in the Niangua River about 1 km downstream from Bennett Spring State Park because growth and fecundity data gathered on hellbenders in those two rivers by Peterson et al. (1983), Topping and Ingersol (1981), and Taber et al. (1975) were compared with data collected in this study.

Population size for each colleciton site was estimated with the Jolly-Seber model (Jolly, 1965; Seber, 1965) as described by Begon (1979) and Caughley (1977). This model assumes that the capture of an individual does not affect the probability of subsequent recaptures and that all previously captured or uncaptured animals are equally catchable (Jolly and Dickson, 1983; Seber, 1982). The form and nature of the statistical tests used to test these assumptions are described by Begon (1979). Population size of site 1 in the Big Piney was also estimated by a frequency of capture method as described by Caughley (1977). This method assumes there are no losses during the study but does not make the equal-catchability assumption. It requires multiple recaptures of individuals.

Length-fecundity relationships were developed by collecting females before the spawning season and counting the eggs to be laid that year: the enlarged yolk-laden ova of both ovaries. The females were transported on ice from the collection site to the laboratory before killing with tricaine or by freezing. Ovaries were removed and placed in $10 \%$ formalin before oocytes were counted. Bartlett's (1949) method as described by Simpson et al. (1960) was used to describe the relationship
between $T L$, the independent variable, and the number of eggs per female, the dependent variable. Davies (1971) provided a Fortran program for this method, and Simpson et al. (1960) described methods for testing for the significance of the slope and for deviations from linearity, for constructing confidence limits for the slope and $y$-intercept, and for comparing two regression lines. If there are errors in the independent variable, least squares regression provides biased, inconsistent estimators for the relationship between the variables unless target values of the independent variable are set or unless severe assumptions can be imposed on the problem (Neter and Wasserman, 1974). Although the least squares equation is still suitable for prediction (Bartlett, 1949), Ricker (1973) has suggested that when both variables are subject to natural variability and measurement error the functional relationship is best described by the geometric mean regression, Wald's regression or the Nair-Bartlett regression. Because Bartlett's method was used succesfully on hellbenders (cf. Peterson et al., 1985) and is easy to apply, this regresion method was used whenever measurement error could exist in the independent variable.

Length-specific growth rates of hellbenders for each population were estimated from recapture data. Mean monthly growth was regressed on initial TL. Growth increments within a single season were rejected to minimize the effect of marking and measurement error on growth. Only initial and final Ths were used to estimate growth of animals recaptured more than once to maintain independence. As described by Van Devender (1978), the length-growth equations were integrated to yield an equation relating age to TL.

Fecundity tables, life tables, and cohort biomass were determined from a sample of females caught within $\pm 2$ months of the breeding season. A stationary age distribution for each population was assumed. Females were aged based on TL. The number of females in the zero age class was estimated by summing the products of the number of females in each age class and the expected number of eggs produced per female (estimated from the length-fecundity relationships). Survivorship was calculated by dividing the zero age class, taken from the fecundity tables, into the number of females alive in each age class. The average age at egg laying or mean generation time was calculated by multiplying the interval midpoint age times the number of zygotes produced per female times the survivorship for each age class and summing all the products (cf. Leslie, 1966). Female cohort biomass was determined from the life tables and the estimated number of females in the collection sites. Average weight at death of the females was estimated from length-weight regressions.

Annual secondary production of females was estimated by the summation-removal method (cf. Waters, 1977) from the female cohort biomass (assuming a stable age distribution). Production : biomass ratios were calculated by multiplying site biomass estimates by the average female proportion of the population and dividing the product into the production estimate. Yearly production was also estimated based on the product of biomass and an estimate of instantaneous growth: 1n (weight of heaviest individual / ovum weight) / hypothesized age of oldest individual.

## RESULTS <br> Population Size

Initial marking mortality. The recapture data (Table 1) were used to estimate population size, recruitment, and survivorship rates by the Jolly-Seber method. To test the assumption that marking does not increase mortality, a chi-square test may be used (Begon, 1979). Of the $r_{i}$ individuals captured and released on day $i$, some will have been caught and marked in samples prior to day $i$ and some will not. If the chances of dying are increased by the initial marking, then the individuals which were marked for the first time on day $i$ will be underrepresented in subsequent samples. Chi-squares and their degrees of freedom may be summed for different sampling periods.

Because methods were the same, the effect of capturing and marking on mortality should be similar for different populations unless the size of individuals being branded is a factor. Therefore the test of whether marking increases mortality was conducted only on the population where individuals were observed to be smallest, the Eleven Point River population (Table 2). Although the chi-square value for the third sampling period is significant at $\alpha=0.1$, the overall value indicates that being caught and marked for the first time did not affect the mortality of Eleven Point hellbenders significantly.

Random sampling. Leslie's (1958) test for random sampling as outlined by Begon (1979) and Caughley (1977) was performed on recapture data for each population except, due to insufficient data, the Eleven Point hellbenders (Table 3). Also the Spring River data are not quite adequate for this test. Of $G$ individuals (should exceed 20) caught on or before day $j$ and on or after day $j+t, g_{i}$ individuals will be caught
Table 1. Summary of the capture-recapture data.

| Study interval | Population | $\begin{aligned} & \text { No. of } \\ & \text { collection } \\ & \text { sites } \end{aligned}$ | No. of individuals captured | No. recaptured at least once | Total recaptures |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 24 May 1980- <br> 21 Aug. 1982 | Spring | 2 | 370 | 92 | 119 |
| $\begin{aligned} & 26 \text { July 1980- } \\ & 4 \text { Sept. } 1982 \end{aligned}$ | Eleven Point | 2 | 211 | 55 | 68 |
| 21 July 1980- 26 Sept. 1982 | Gasconade | 1 | 293 | 126 | 240 |
| 31 May 1980- <br> 24 July 1982 | Big Piney | 2 | 334 | 135 | 227 |
| Totals |  | 7 | 1208 | 408 | 654 |

Table 2. Chi-square test of whether initial marking has an effect on the mortality of Eleven Point hellbenders.

| Sample | 2 | df | $\boldsymbol{\alpha}=0.1$ | $\boldsymbol{\alpha}=0.5$ |
| :---: | :---: | :---: | :---: | :---: |
| 2 | 0.39 | 1 | ns | ns |
| 3 | 2.95 | 1 | $*$ | ns |
| 4 | 2.09 | 1 | ns | ns |
| overall | 5.43 | 3 | ns | ns |

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Table 3. Leslie's (1958) approximate chi-square test comparing the expected variance of the observed recapture frequencies.

| Population | No. known alive between time i and $i+j$ | No. of samples between $i$ and $i+j$ | No. of individuals captured (0-3 times) between $i$ and $i+j$ |  |  |  | 2 | $\boldsymbol{\alpha}=0.1$ | $\boldsymbol{\alpha}=0.05$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | 0 | 1 | 2 | 3 |  |  |  |
| Spring | 15 | 3 | 10 | 3 | 2 | 0 | 19.72 | ns | ns |
| Eleven Point | $4^{\text {a }}$ | 3 |  |  |  |  |  |  |  |
| Gasconade | 35 | 3 | 12 | 11 | 11 | 2 | 47.39 | * | ns |
| Big Piney | 22 | 3 | 5 | 13 | 5 | 1 | 20.22 | ns | ns |

[^0]on day $i$ (should be at least three day i) between $j$ and $j+t$. If the catchability is constant, the recapture frequencies will form a binomial distribution. This hypothesis can be tested by comparing the observed variance of recapture frequencies with the expected binomial variance. This approximate chi-square test is run with $\Sigma g_{i}-1$ degrees of freedom. There is no indication that sampling was not random in the three populations tested at the 0.05 significance level, although the Gasconade value exceeds the critical value at the 0.1 level.

Roff (1973) argued that Leslie's test is incapable of distinguishing whether subgroups, such as sex and age classes, differ in catchability. Thus, because data are insufficient to obtain reliable density estimates directly on different subgroups, tests of whether male and female hellbenders were equally catchable and/or had similar survival rates were performed on each population (Table 4). Similar tests were made on three life stages: animals were classified as juveniles-young adults, adults, or old adults based on TL (Table 5). These tests can be used to test for differences in capture probabilities alone by assuming a constant survival probability. Alternatively, if the capture probabilities are assumed constant, the test is for constancy of survival probability (Begon, 1979). These tests can be made with single-mark-recapture data or multiple-mark-recapture data. Considering the longevity of hellbenders (Peterson et al., 1983; Taber et al., 1975), the single-mark-recapture data (only for samples within the same year) should provide information on equal catchability nearly exclusive of survivorship differences in the subgroups. Only
Table 4. Chi-square tests of whether the survival-rates and/or the catchabilities of males and females are the same. The test is repeated for each sample period, and the chi-square values and degrees of freedom are summed as suggested by Begon capture-recapture tests to decrease survivorship effects.

| Population | Multiple capture-recapture |  |  | Single capture-recapture |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\chi^{2}$ | df | Significance | $\chi^{2}$ | df | Significance |
| Spring | 0.51 | 1 | ns | 0.04 | 1 | ns |
|  | 1.76 | 1 | ns | 0.18 | 1 | ns |
|  | 0.53 | 1 | ns | 0.22 | 2 | ns |
|  | 1.07 | 1 | ns |  |  |  |
|  | 3.87 | 4 | ns |  |  |  |
| Eleven Point | 2.35 | 1 | ns | 1.81 | 1 | ns |
|  | 0.02 | 1 | ns |  |  |  |
|  | $\overline{2.37}$ | - | ns |  |  |  |

*, **, *** $\mathrm{P}<0.1,0.05,0.01$ respectively
Table 4 cont. Chi-square tests of whether the survival-rates and/or the catchabilities of males and females are the same. The test is repeated for each sample period, and the chi-square values and degrees of freedom are summed as suggested by Begon (1979). Only samples within the same calendar year are considered in
the single capture-recapture tests to decrease survivorship effects.
Multiple capture-recapture Single capture-recapture

*, **, *** $\mathrm{P}<0.1,0.05,0.01$ respectively
Table 5. Chi-square tests of whether the survival-rates and/or the catchabilities
of juveniles-young adults, adults, and old adults are the same. The test is
repeated for each sample period, and the chi-square values and degrees of freedom
are summed as suggested by Begon (1979). Only samples within the same calendar year are considered in the single capture-recapture tests to decrease survivorship effects.

| Population | Multiple capture-recapture |  |  | Single capture-recapture |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $x^{2}$ | df | Significance | $x^{2}$ | df | Significance |
| Spring | 3.48 | 2 | ns | 5.09 | 2 | * |
|  | 3.39 | 2 | ns | 0.94 | 2 | ns |
|  | 1.68 | 2 | ns | 4.77 | 2 | * |
|  | 1.45 | 2 | ns | 3.38 | 2 | ns |
|  | 1.15 | 2 | ns | 0.59 | 2 | ns |
|  | 1.36 | 2 | ns | 0.41 | 2 | ns |
|  | 0.60 | 2 | ns | 15.18 | $\overline{12}$ | ns |
|  | 0.96 | 2 | ns |  |  |  |
|  | 14.07 | 16 | ns |  |  |  |

*, **, *** $\mathrm{P}<0.1,0.05,0.01$ respectively
Table 5 cont. Chi-square tests of whether the survival-rates and/or the catchabilities of juveniles-young adults, adults, and old adults are the same. The
of freedom are summed as suggested by Begon (1979). Only samples within the
same calendar year are considered in the single capture-recapture tests to decrease survivorship effects.

| Population | Multiple capture-recapture |  |  | Single capture-recapture |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\chi^{2}$ |  | Significance | $x^{2}$ | df | Significance |
| Eleven Point | 3.30 | 2 | ns | 2.00 | 2 | ns |
|  | 0.39 | 2 | ns | 1.01 | 2 | ns |
|  | 1.82 | 2 | ns | 3.01 | 4 | ns |
|  | 0.13 | 2 | ns |  |  |  |
|  | $\overline{5.64}$ | 8 | ns |  |  |  |

[^1]Table 5 cont. Chi-square tests of whether the survival-rates and/or the catchabilities of juveniles-young adults, adults, and old adults are the same. The test is repeated for each sample period, and the chi-square values and degree decrease survivorship effects.

| Population | Multiple capture-recapture |  |  | Single capture-recapture |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\chi^{2}$ | df | Significance | $\chi^{2}$ | df | Significance |
| Gasconade | 5.46 | 2 | * | 3.61 | 2 | ns |
|  | 2.14 | 2 | ns | 8.33 | 2 | ** |
|  | 5.53 | 2 | * | 2.33 | 2 | ns |
|  | 0.96 | 2 | ns | 0.80 | 2 | ns |
|  | 0.38 | 2 | ns | 2.34 | 2 | ns |
|  | 1.25 | 2 | ns | 1.36 | 2 | ns |
|  | 4.72 | 2 | * | 18.77 | 12 | * |
|  | 1.44 | 2 | ns |  |  |  |
|  | 21.88 | 16 | ns |  |  |  |

[^2]Table 5 cont. Chi-square tests of whether the survival-rates and/or the catchabilities of juveniles-young adults, adults, and old adults are the same. The test is repeated for each sample period, and the chi-square values and degrees of freedom are summed as suggested by Begon (1979). Only samples within the same calendar year are considered in the single capture-recapture tests to decrease survivorship effects.

| Population | Multiple capture-recapture |  |  | Single capture-recapture |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $x^{2}$ |  | Significance | $\chi^{2}$ | df | Significance |
| Big Piney | 3.37 | 2 | ns | 2.90 | 2 | ns |
|  | 3.43 | 2 | ns | 6.61 | 2 | ** |
|  | 1.12 | 2 | ns | 1.28 | 2 | ns |
|  | 4.72 | 2 | * | 3.78 | 2 | ns |
|  | 2.04 | 2 | ns | 7.99 | 2 | ** |
|  | 3.29 | 2 | ns | 11.06 | 2 | *** |
|  | 7.28 | 2 | ** | 33.62 | 12 | *** |
|  | 11.06 | 2 | *** |  |  |  |
|  | 36.31 | $\overline{16}$ | *** |  |  |  |

[^3]```
in the Big Piney population did males and females differ significantly
in survivorship and/or catchability. Also only in the Big Piney were
individuals of different life stages significantly different with
regard to catchability.
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Population size, recruitment, and survival rates. Besides the estimates of population size provided by the Jolly-Seber model (Table 6), estimates for the first and last collection periods were made by the method described by Caughley (1977). Due to difficulties with equal catchability of subgroups in the Big Piney, population size for site 1 , where there were multiple recaptures of individuals, was also estimated by a frequency of capture model that does not require random sampling but assumes no losses. Goodness-of-fit tests of the frequency of capture of individuals to a zero-truncated Poisson, negative binomial, and geometric distribution were made by chi-square tests. The best fit of the observed frequency of capture was to the geometric distribution $\left(x^{2}=4.26, \mathrm{P}>0.10\right)$ yielding a population estimate of 806 hellbenders. The average estimate for site 1 by the Jolly-Seber method was 161 hellbenders.

The recruitment estimates ( $A_{i}$ ) include animals that have entered the population through natality or immigration. The average of the $A_{i}$ for each collection site indicates little increase or decrease in numbers during the study period. The observed rate of increase can be estimated by the slope of the least squares regression of the natural logarithm of the population estimate versus time (Caughley, 1966). For all collection sites but site 1 of the Eleven Point (Fig. 1), which has a significant negative slope (t=3.96, $\mathrm{P}<0.05$ ), the slopes are not significantly different from zero (all $t<1.96$ ).
Table 6. Jolly-Seber estimates of population size (N), number of individuals entering population between time $i$ and $i+1$ (A), probabilities that an individual present at time $i$ will still be present at time $i+1(P)$, and their standard errors (SE).
Monthly survival rates (MSR) are the nth roots of the corresponding $P_{i}$ 's, where $n=$ the number of months between time $i$ and $i+1$.

| Population | $\begin{aligned} & \text { Interval } \\ & \text { (i to } \mathrm{i}+1 \text { ) } \end{aligned}$ | Interval length (months) | N | SE | A | SE | P | SE | MSR |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Spring | 0-1 | - | $459.4{ }^{\text {a }}$ | - | - | - | - | - | - |
| Site 1 | 1-2 | 0.6 | 463.8 | 97.8 | -177.7 | 500.8 | 1.213 | 0.276 | 1.380 |
|  | 2-3 | 0.6 | 384.8 | 79.7 | 409.2 | 1252.4 | 1.361 | 0.377 | 1.671 |
|  | 3-4 | 1.2 | 933.0 | 226.2 | -187.2 | 541.4 | 0.638 | 0.186 | 0.688 |
|  | 4-5 | 3.0 | 408.0 | 95.9 | -88.6 | 365.2 | 0.772 | 0.302 | 0.917 |
|  | 5-6 | 8.0 | 226.4 | 78.8 | 202.2 | 1077.5 | 1.096 | 0.581 | 1.012 |
|  | 6-7 | 0.8 | 450.4 | 188.8 | -11.2 | 89.1 | 0.534 | 0.337 | 0.456 |
|  | 7-8 | 1.4 | 229.3 | 116.8 | - | - | - | - | - |
|  | 8-9 | 11.3 | $228.9{ }^{\text {a }}$ | - | - | - | - | - | - |
| $\overline{\mathrm{x}}$ and $\mathrm{SE}_{\overline{\mathrm{x}}}$ |  |  | 420.4 | 72.6 | 24.4 | 96.5 | 0.936 | 0.137 | 1.021 |
| Site 2 | 0-1 | - | $120.3{ }^{\text {a }}$ | - | - | - | - | - | - |
|  | 1-2 | 0.7 | 121.3 | 113.2 | - | - | - | - | - |
|  | 2-3 | 0.6 | 55.2 | 24.9 | 22.6 | 64.5 | 0.808 | 0.512 | 0.701 |
|  | 3-4 | 1.3 | 67.2 | 34.7 | - | - | - | - | - |
|  | 4-5 | 2.0 | $67.3{ }^{\text {a }}$ | - | - | - | - | - | - |
| $\overline{\mathrm{X}}$ and $\mathrm{SE}_{\overline{\mathrm{X}}}$ |  |  | 86.3 | 14.3 |  |  |  |  |  |

$\mathrm{a}_{\text {Estimate }}$ calculated by method described by Caughley (1977).
Table 6 cont. Jolly-Seber estimates of population size (N), number of individuals entering population between:times $i$ and $i+1$ (A), probabilities that an individual present at time $i$ will still be present at time $i+1(P)$, and their standard errors (SE). Monthly survival rates (MSR) are the nth roots of the corresponding $P_{i}$ 's, where $n=$ the number of months between time $i$ and $i+1$.

| Population | $\begin{aligned} & \text { Interval } \\ & \text { (i to } i+1 \text { ) } \end{aligned}$ | Interval <br> length <br> (months) | N | SE | A | SE | P | SE | MSR |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Eleven Point | 0-1 | - | $170.7^{\text {a }}$ | - | - | - | - | - | - |
| Site 1 | 1-2 | 1.8 | 171.9 | 54.5 | 0.7 | 49.4 | 0.676 | 0.246 | 0.804 |
|  | 2-3 | 10.0 | 117.0 | 33.5 | -3.7 | 15.5 | 0.451 | 0.190 | 0.923 |
|  | 3-4 | 3.0 | 49.1 | 19.0 | - | - | - | - | - |
|  | 4-5 | 9.0 | $48.9{ }^{\text {a }}$ | - | - | - | - | - | - |
| $\overline{\mathrm{X}}$ and $\mathrm{SE}_{\overline{\mathrm{x}}}$ |  |  | 111.5 | 27.4 | $-1.5$ | 2.2 | 0.564 | 0.112 | 0.864 |
| Site 2 | 0-1 | - | $110.2^{\text {a }}$ | - | - | - | - | - | - |
|  | 1-2 | 1.9 | 111.1 | 35.8 | -5.6 | 24.6 | 0.702 | 0.250 | 0.830 |
|  | 2-3 | 10.0 | 72.4 | 25.0 | - | - | - | - | - |
|  | 3-4 | 13.2 | $72.3{ }^{\text {a }}$ | - | - | - | - | - | - |
| $\overline{\mathrm{X}}$ and $\mathrm{SE}_{\overline{\mathrm{x}}}$ |  |  | 91.5 | 11.1 |  |  |  |  |  |

[^4]Table 6 cont. Jolly-Seber estimates of population size (N), number of individuals entering population between times $i$ and $i+1$ (A), probabilities that an individual present at time $i$ will still be present at time $i+1(P)$, and their standard errors (SE). Monthly survival rates (MSR) are the nth roots of the corresponding $P_{i}$ ' $s$, where $n=$ the number of months between time $i$ and $i+1$.

| Population | $\begin{aligned} & \text { Interval } \\ & \text { (i to } i+1 \text { ) } \end{aligned}$ | Interval length (months) | N | SE | A | SE | P | SE | MSR |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Gasconade | 0-1 | - | $96.8{ }^{\text {a }}$ | - | - | - | - | - | - |
| Site 1 | 1-2 | 0.6 | 100.4 | 6.3 | 88.8 | 73.0 | 1.021 | 0.102 | 1.035 |
|  | 2-3 | 0.7 | 191.2 | 16.9 | 196.1 | 377.8 | 1.100 | 0.182 | 1.146 |
|  | 3-4 | 1.9 | 406.5 | 63.8 | -66.7 | 155.0 | 0.707 | 0.158 | 0.833 |
|  | 4-5 | 1.6 | 220.5 | 41.4 | 80.2 | 221.8 | 0.720 | 0.214 | 0.814 |
|  | 5-6 | 9.2 | 238.9 | 57.3 | -33.2 | 125.0 | 1.112 | 0.290 | 1.012 |
|  | 6-7 | 0.5 | 232.5 | 28.2 | 42.0 | 105.9 | 1.118 | 0.231 | 1.250 |
|  | 7-8 | 0.8 | 302.0 | 54.8 | -6.9 | 28.0 | 0.594 | 0.176 | 0.521 |
|  | 8-9 | 10.0 | 164.7 | 40.8 | - | - | - | - | - |
|  | 9-10 | 2.0 | $162.7{ }^{\text {a }}$ | - | - | - | - | - | - |
| $\overline{\mathrm{x}}$ and $\mathrm{SE}_{\overline{\mathrm{x}}}$ |  |  | 211.6 | 29.4 | 42.9 | 33.5 | 0.910 | 0.086 | 0.944 |

astimate calculated by method described by caughley (1977).
Table 6 cont. Jolly-Seber estimates of population size (N), number of individuals entering population between times $i$ and $i+1$ (A), probabilities that an individual present at time $i$ will still be present at time $i+1$ ( $p$ ), and their standard errors (SE). Monthly survival rates (MSR) are the nth roots of the corresponding $P_{i}$ 's, where $n=$ the number of months between time $i$ and $i+1$.

| Population | $\begin{aligned} & \text { Interval } \\ & \text { (i to } i+1 \text { ) } \end{aligned}$ | $\begin{aligned} & \text { Interval } \\ & \text { length } \\ & \text { (months) } \end{aligned}$ | N | SE | A | SE | P | SE | MSR |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Big Piney | 0-1 | - | $160.3^{\text {a }}$ | - | - | - | - | - | - |
|  | 1-2 | 0.2 | 154.1 | 16.8 | 64.7 | 84.3 | 0.800 | 0.122 | 0.328 |
|  | 2-3 | 1.5 | 187.9 | 24.9 | 17.0 | 36.3 | 0.842 | 0.152 | 0.892 |
|  | 3-4 | 0.6 | 175.2 | 26.3 | 34.9 | 65.9 | 0.669 | 0.150 | 0.512 |
|  | 4-5 | 1.3 | 152.1 | 29.0 | 70.2 | 23.0 | 1.191 | 0.401 | 1.144 |
|  | 5-6 | 9.5 | 251.3 | 74.1 | 6.4 | 29.1 | 0.459 | 0.138 | 0.921 |
|  | 6-7 | 2.1 | 121.8 | 17.8 | -11.8 | 42.0 | 1.100 | 0.347 | 1.046 |
|  | 7-8 | 10.2 | 122.2 | 27.0 | - | - | - | - | - |
|  | 8-9 | 0.6 | $122.2{ }^{\text {a }}$ | - | - | - | - | - | - |
| $\overline{\mathrm{x}}$ and $\mathrm{SE}_{\overline{\mathrm{X}}}$ |  |  | 160.8 | 13.8 | 30.2 | 13.3 | 0.844 | 0.111 | 0.807 |
| Site 2 | 0-1 | - | $69.6{ }^{\text {a }}$ | - | - | - | - | - | - |
|  | 1-2 | 0.6 | 71.4 | 19.7 | -5.2 | 18.0 | 0.720 | 0.258 | 0.578 |
|  | 2-3 | 10.7 | 46.2 | 10.1 | 68.2 | 496.6 | 1.250 | 0.116 | 1.021 |
|  | 3-4 | 1.9 | 126.0 | 114.9 | - | - | - | - | - |
|  | 4-5 | 10.6 | $126.4{ }^{\text {a }}$ | - | - | - | - | - | - |
| $\overline{\mathrm{x}}$ and $\mathrm{SE}_{\overline{\mathrm{x}}}$ |  |  | 87.9 | 16.2 | 31.5 | 36.7 | 0.985 | 0.265 | 0.800 |

[^5]Figure 1. Natural logarithm (LN) of the estimated number of
hellbenders during the study period for each
collection site from Table 6.

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```
    Estimates of probability of survival ( ( }\mp@subsup{\textrm{i}}{\mathbf{i}}{\prime}\mathrm{ ) from time i to i+1
do not distinguish between actual mortality and emigration from the
site, Due to sampling error, the }\mp@subsup{P}{i}{}\mathrm{ may exceed one. Monthly survival
rate (MSR) was calculated as }\mp@subsup{P}{i}{}\mp@subsup{}{}{\mathrm{ -interval length as described by}
Tilley (1980). A comparison of MSRs for all sites of the Big Piney,
Gasconade, and Spring River populations was made by ANOVA after a test
of equal variance was made using Hartley's test (H=4.19, r=3, n=7.3,
P>0.05) as outlined by Neter and Wasserman (1974). No significant
difference ( }\textrm{F}=0.143,P>0.05) in mean MSRs was indicated. Th
Eleven Point population was excluded from the comparison due to the
paucity of estimates. An average MSR of 0.8974 (SD=0.3014) was
calculated from the pooled MSRs, including those of the Eleven Point
sites. The annual survivorship rate was then estimated as 0.2728
for hellbenders from all collection sites of all populations.
    Densities and biomass. Relative abundance was estimated in the
different populations by determining the number of hellbenders caught
per time (Table 7). Timed samples were made only when visibility was
good, the river was not in flood, and temperatures were not extremely
cold. From the average wet weight of an individual caught from each
collection site (Table 8) and the average population estimate (Table 6),
site biomass was calculated (Table 9). There is a significant
correlation between absolute density of hellbenders (no./100 m}\mp@subsup{}{}{2}\mathrm{ ) and
mean relative density (catch/hr) estimates (Table 9), except for the
Gasconade population (r=0.867, P<0.05, exclusive of the Gasconade
estimates).
```

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Table 7. Average number of hellbenders caught per hour.

| Population | Site | No. of timed samples | Average <br> catch <br> time <br> (hours) | Average no. of hellbenders caught per hour | SD | Range |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Spring | 1 | 6 | 2.47 | 9.02 | 1.98 | 6.43-12.00 |
|  | 2 | 3 | 2.17 | 3.81 | 1.67 | 2.40-5.65 |
| Eleven | 1 | 3 | 1.14 | 8.40 | 3.12 | 6.46-12.00 |
| Point | 2 | 2 | 0.83 | 8.75 | 1.77 | $7.50-10.00$ |
| Gasconade | 1 | 3 | 1.69 | 13.21 | 3.12 | 11.08-16.80 |
| Big Piney | 1 | 3 | 1.28 | 7.64 | 2.55 | 5.05-10.15 |
|  | 2 | 2 | 1.00 | 12.75 | 4.24 | $9.75-15.75$ |

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Table 9. Density and biomass of hellbenders

| Population | Site | $\begin{aligned} & \text { Area } \\ & \left(\mathrm{m}^{2}\right) \end{aligned}$ | No. of estimates | Average estimate | SD | No. per$100 \mathrm{~m}^{2}$ | Site biomass |  | Catch/hour |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  | kg | kg/ha | mean | SD |
| Spring | 1 | 10,300 | 9 | 420 | 218 | 4.1 | 407.7 | 396 | 9.02 | 1.98 |
|  | 2 | 9,000 | 5 | 86 | 32 | 1.0 | 66.2 | 74 | 3.81 | 1.67 |
| Eleven Point | 1 | 2,600 | 5 | 112 | 61 | 4.3 | 24.0 | 92 | 8.40 | 3.12 |
|  | 2 | 1,500 | 4 | 92 | 22 | 6.1 | 31.3 | 209 | 8.75 | 1.77 |
| Gasconade | 1 | 21,400 | 10 | 212 | 93 | 1.0 | 126.7 | 59 | 13.21 | 3.12 |
| Big Piney | 1 | 7,000 | 9 | 161 | 42 | 2.3 | 78.4 | 112 | 7.64 | 2.55 |
|  | 2 | 1,400 | 5 | 88 | 36 | 6.3 | 46.8 | 334 | 12.75 | 4.24 |

## Population Composition

Except in the Spring River where there was an unequal adult sex ratio, adult females appeared to be more common than adult males, at least from late July into November when hellbenders can be sexed externally (Table 10). In the Gasconade and Spring Rivers the populations were dominated by large, long-lived adults. Only one juvenile in the latter river and two in the former were captured during late summer to fall during three years, but $200^{+}$adults were captured in each river during those samples. Only in the Eleven Point was a substantial proportion (25\%) of the population composed of immature individuals. Further, of the approximately 1200 hellbenders marked during this study, only one larval or gilled specimen was observed. This individual from the Big Piney measured 77 mm TL and was approximately 10 months old from time of spawning.

## Fecundity

Paired comparison of ovary oocyte number. All externally sexed females were ripe but one. The number of developing eggs in the right ovary was compared with those of the left ovary for each population by a two-way ANOVA without replication (Sokal and Rolf, 1973) after no significant difference was found in the variance of ova number in right and left ovaries (Table 11). A significantly greater number of ova in the right ovary was indicated for the Eleven Point and Gasconade animals, but no significant difference between ovaries in the Spring and Big Piney females.

Size at maturity. Size at sexual maturity seems to differ among populations. From external identification Eleven Point females mature at about 300 mm TL while Big Piney females do not mature
Table 10. Number and percent of juveniles, adult males, and adult females in samples where all adults were sexed externally and chi-square tests of whether the ratio of adult males and females is significally different from l:1 for each sample and for the total number of individuals (adult males and females) in those samples (excludes recaptures).

| Date | Juveniles |  | Males |  | Females |  | $\chi^{2}$ | Total no. of individuals in samples |  | $x^{2}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | No. | \% | No. | \% | No. | \% |  | M | F |  |
| 9-10 Aug. 1980 | 0 | 0.0 | 40 | 44.4 | 50 | 55.6 | $1.11{ }^{\text {ns }}$ | 128 | 119 | $0.33{ }^{\text {ns }}$ |
| 8-9 Nov. 1980 | 0 | 0.0 | 42 | 56.0 | 33 | 44.0 | $1.08{ }^{\text {ns }}$ |  |  |  |
| 31 July 1981 | 0 | 0.0 | 17 | 54.8 | 14 | 45.2 | $0.29{ }^{\text {ns }}$ |  |  |  |
| 12 Sept. 1981 | 0 | 0.0 | 22 | 51.2 | 21 | 48.8 | $0.02{ }^{\text {ns }}$ |  |  |  |
| 21 Aug. 1982 | 1 | 2.6 | 17 | 43.6 | 21 | 53.8 | 0.42 ns |  |  |  |
|  |  | 0.5 | $\overline{\mathrm{x}}=$ | 50.0 |  | 49.5 | 2.93 ns |  |  |  |

[^6]Table 10 cont. Number and percent of juveniles, adult males, and adult females in samples where all adults were sexed externally and chi-square tests of whether the ratio of adult males and females is significantly different from $1: 1$ for each sample and for the total number of individuals (adult males and females) in those samples (excludes recaptures).

| Date | Juveniles |  | Eleven Point River |  |  |  | $x^{2}$ | Total no. of individuals in samples |  | $x^{2}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Males |  | Females |  |  |  |  |  |
|  | No. | \% | No. | \% | No. | \% |  | M | F |  |
| 26-27 July 1980 | 23 | 26.4 | 23 | 26.4 | 41 | 47.1 | 5.06** | 54 | 86 | 7.31*** |
| 20-21 Sept. 1980 | 19 | 25.3 | 24 | 32.0 |  | 42.7 | $1.14{ }^{\text {ns }}$ |  |  |  |
| 18 Oct. 1981 | 9 | 47.4 | 2 | 10.5 | 8 | 42.1 | 3.60* |  |  |  |
| 4 Sept. 1982 | 0 | 0.0 | 9 | 31.0 |  | 69.0 | 4.17**, |  |  |  |
|  |  | 24.8 | $\overline{\mathrm{x}}=$ | 25.0 | $\overline{\mathrm{x}}=$ | 50.2 | 13.98*** |  |  |  |

*, **, *** $\mathrm{P}<0.1,0.05,0.01$ respectively

Table 10 cont. Number and percent of juveniles, adult males, and adult females in samples where all adults were sexed externally and chi-square tests of whether the ratio of adult males and females is significantly different from $1: 1$ for each sample and for the total

| Date | Gasconade River |  |  |  |  |  |  | Total no. of individuals in samples |  | $\chi^{2}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Juveniles |  | Males |  | Females |  | $\chi^{2}$ |  |  |  |
|  | No. | \% | No. | \% | No. | \% |  | M | F |  |
| 28 Sept. 1980 | 0 | 0.0 | 22 | 34.4 | 42 | 65.6 | $6.25 * *$ | 95 | 131 | 5.74** |
| 15 Nov. 1980 | 0 | 0.0 | 5 | 20.0 | 20 | 80.0 | 9.00*** |  |  |  |
| 5 Sept. 1981 | 1 | 1.8 | 21 | 36.8 | 35 | 61.4 | 3.50* |  |  |  |
| 28 Sept. 1981 | 1 | 1.4 | 36 | 50.0 | 35 | 48.6 | $0.01{ }^{\text {ns }}$ |  |  |  |
| 27 July 1982 | 0 | 0.0 | 28 | 39.4 | 43 | 60.6 | 3.17* |  |  |  |
| 26 Sept. 1982 | 0 | 0.0 | 12 | 35.3 | 22 | 64.7 | 2.94* |  |  |  |
|  |  | 0.5 | $\overline{\mathrm{x}}$ | 36.0 | X | 63.5 | 24.87*** |  |  |  |

*, **, *** $\mathrm{P}<0.1,0.05,0.01$ respectively
Table 10 cont. Number and percent of juveniles, adult males, and adult females in samples where all adults were sexed externally and chi-square tests of whether the ratio of adult males and females is significantly different from $1: 1$ for each sample and for the total number of individuals (adult males and females) in those samples (excludes recaptures).

*, **, *** $\mathrm{P}<0.1,0.05,0.01$ respectively
and left ovaries by analysis of variance.

| Population | N | Left ovary |  | Right ovary |  | F value |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Mean no. of ova | SD | Mean no. of ova | SD | Variance ratio | Paired comparison |
| Spring | 15 | 283.3 | 71.1 | 293.2 | 103.5 | $2.119^{\mathrm{ns}}$ | $0.428{ }^{\text {ns }}$ |
| Eleven Point | 15 | 135.8 | 39.9 | 153.1 | 41.5 | $1.082^{\mathrm{ns}}$ | 15.096** |
| Gasconade | 16 | 179.9 | 54.0 | 194.1 | 64.7 | $1.436^{\text {ns }}$ | 4.605* |
| Big Piney | 16 | 156.9 | 54.8 | 162.9 | 51.2 | $1.146^{\text {ns }}$ | $0.953{ }^{\text {ns }}$ |

*, **, *** $\mathrm{P}<0.05,0.01,0.001$ respectively

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until a size of about 370 mm . Due to the few number of juveniles, size at maturity could not be determined for Spring River and Gasconade River females. All females dissected from the Eleven Point were mature. The smallest of these was 323 mm TL . One 353 mm female from the Big Piney was immature, the next smallest individual dissected, 387 mm TL , was mature. No female less than 400 mm was dissected from either the Spring or Gasconade Rivers.

TL-fecundity relationships. Fecundity is positively related to female length in all populations (Table 12 and Fig. 2). The Gasconade females and the Big Piney females exhibit similar TL-fecundity relationships as do the two C. a. bishopi populations (Table 13); however, the latter two were not pooled in further comparisons because of the disparity in size of Spring River and Eleven Point females. The TL-fecundity regressions calculated for North Fork and Niangua River females by Topping and Ingersol (1981) were included in the comparisons. The females of the three C. a. bishopi populations produce more eggs at any specific size than the females of the three C. a. alleganiensis populations. All possible comparisons were not made because of the increasing possibility of error due to chance alone with each further test. Moreover, further comparisons would provide little more information about the relationships between the regressions, especially because sample sizes are similar.

Spawning. Thirteen of 15 Big Piney females killed 19 September 1981 were spent. On 26 September 1981, egg strings were protruding from the cloaca of a few captured Gasconade females, and an egg string was found under a rock. Seven Spring River females were
Table 12. Regression of fecundity (F) on total length (TL) using Bartlett's (1949) method intercept, respectively, are shown in parentheses.

| Population | N | $\begin{gathered} \text { Mean } \\ \text { TL } \\ (\mathrm{mm}) \end{gathered}$ | SD | Mean clutch size | SD | Regression equation | $r^{2}$ | t value |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  | nonlinear component | slope differing from zero |
| Spring | 15 | 480 | 51 | 576 |  | $\begin{gathered} F=2.3374 \mathrm{TL}-545.2915 \\ 0.85-3.64)(-608--482) \end{gathered}$ | 0.614 | $0.398^{\mathrm{ns}}$ | 3.223** |
| Eleven Point | 15 | 365 | 33 | 289 |  | $\begin{gathered} F=1.8405 T L-382.9678 \\ 0.64-2.87) \quad(-404--362) \end{gathered}$ | 0.673 | $0.807^{\text {ns }}$ | 2.972* |
| Gasconade | 16 | 450 | 36 | 374 |  | $\begin{gathered} F=2.4066 \mathrm{TL}-709.0420 \\ 1.03-3.54) \quad(-746--672) \end{gathered}$ | 0.691 | $0.714^{\text {ns }}$ | 3.321** |
| Big Piney | 16 | 429 | 31 | 320 |  | $\begin{gathered} \mathrm{F}=2.7258 \mathrm{TL}-849.8760 \\ 1.47-4.10)(-887-813) \end{gathered}$ | 0.618 | $0.479^{\text {ns }}$ | 4.583*** |

[^7]```
Figure 2. TL-fecundity regressions. C. a. bishopi populations:
Spring River (S,0), Eleven Point (EP,0). C. a.
alleganiensis populations: Gasconade (G,*), Big
Piney (BP,t). Line lengths represent the approxi-
mate range of adult female sizes for each population.
```

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Table 13. T-tests of the significance of the difference between slopes and y-intercepts
of the regressions of fecundity on total length as described by Simpson et al. (1960).
Regressions for Niangua and North Fork River populations are from Topping and Ingersol
(1981).

| Test | N | $\qquad$ | $\frac{\text { Slopes }}{t}$ | $\frac{\text { Elevations }}{\mathrm{t}}$ | Conclusions |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | variances | slopes | elevations |
| Gasconade vs. Big Piney | 16;16 | $1.030^{\mathrm{ns}}$ | $-0.012^{\mathrm{ns}}$ | $1.889^{\mathrm{ns}}$ | equal | equal | equal |
| Gasconade and Big Piney vs. Niangua | 32;9 | $2.779^{\mathrm{ns}}$ | $0.091{ }^{\text {ns }}$ | -6.199*** | equal | equal | different |
| North Fork vs. Spring | 12;15 | $1.110^{\mathrm{ns}}$ | $-0.032{ }^{\text {ns }}$ | -3.302* | equal | equal | different |
| North Fork vs. Gasconade and Big Piney | 12;32 | $1.247^{\mathrm{ns}}$ | $-0.058{ }^{\text {ns }}$ | -17.342*** | equal | equal | different |
| North Fork vs. Eleven Point | 12;15 | $1.073{ }^{\text {ns }}$ | $-0.243^{\text {ns }}$ | 24.048*** | equal | equal | different |
| ```Spring vs.a Gasconade and Big Piney``` | 15;32 | $1.998{ }^{\text {ns }}$ | $-0.109^{\text {ns }}$ | -44.483*** | equal | equal | different |
| Spring vs. ${ }^{\text {a }}$ Eleven Point | 15;15 | $1.719^{\mathrm{ns}}$ | $0.025^{\text {ns }}$ | $0.736{ }^{\text {ns }}$ | equal | equal | equal |

${ }^{\text {Comparison }}$ made after a double natural logarithmic transformation to equalize variances
$\star, \star *, \star \star * \mathrm{P}<0.05,0.01,0.001$ respectively
collected 12 December 1984; all were ripe. Two males were producing milt.
Stomach Contents
Crayfish remains were the most frequently found items in stomachs,
except for the females collected near the height of the Gasconade
breeding season (Table 14). In 10 of the 13 Gasconade females killed
that day, hellbender eggs were found in the stomach, generally in
large numbers, though they could not be counted. Only four of the
15 Big Piney females contained eggs in the stomach at or after
spawning; five of the 15 contained crayfish remains. Food items regur-
gitated in capture and transport were not considered. It is unknown
how widespread regurgitation was in capture, but it was minimal during
transport on ice.

Growth and Size Comparisons
Growth. Growth of the sexes was pooled in the Spring and Eleven Point populations because of insufficient data to produce significant regressions for adult males and females separately. However, in the Gasconade and Big Piney populations, adult males and females could be treated separately (Fig. 3). The slopes of all regressions shown of growth versus length are negative and significantly different from zero (Table 15). Because of nonlinearity, a logarithmic transformation was performed on the TLs of Gasconade males. There was a significant difference in the $y$-intercepts of Gasconade males and females $(t=-6.22, P<0.001)$ but not slopes $(t=0.03, P>0.05)$. Similarly, females are generally longer than males in the Big Piney River; slopes are not significantly different $(t=0.07, P>0.05)$, but elevations are significantly different ( $\mathrm{t}=7.52, \mathrm{P}<0.001$ ).
Table 14. Contents of stomachs of hellbenders killed for information on fecundity.

| Population | No. of stomachs examined before/ just after spawning | No. of stomachs with crayfish remains before/ just after spawning | Average no. of pairs of chelae per stomach (SD) before spawning | No. of stomachs with fish remains | No. of stomachs with hellbender eggs | No. of stomachs with pebbles/ bits of foliage | No. of stomachs empty |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Spring | 15/0 | 15/- | 2.7(1.5) | 1 | - | 3/1 | 0 |
| Eleven Point | 15/0 | 14/- | 1.9(1.1) | 2 | - | 1/0 | 1 |
| Gasconade | 11/13 | 6/3 | 0.6(0.7) | 0 | 10 | 0/1 | 6 |
| Big PIney | 15/15 | 7/5 | 0.9(1.3) | 1 | 4 | 4/1 | 13 |

```
Figure 3. Length-growth regressions for Spring River,
    Eleven Point, Gasconade, and Big Piney hellbenders.
    Only the Eleven Point regression includes juvenile
    growth. Sexes are pooled for the Spring River and
    Eleven Point regressions.
```

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Table 15. Regression of growth (GR) on total length (TL) using Bartlett's (1949)
method as described by Simpson et al. (1960). The 95\% confidence limits for the
slope and $y$-intercept, respectively, are shown in parentheses.
-

$$
\text { Population Sex } N \quad \text { Regression equation } \quad r^{2} \quad \begin{aligned}
& \text { nonlinear } \begin{array}{l}
\text { differing } \\
\text { component } \\
\text { from zero }
\end{array} \\
& \hline
\end{aligned}
$$

| Spring | Both | 44 | $\begin{gathered} \mathrm{GR}=-0.0079 \mathrm{TL}+4.4134 \\ (-.0038--.0159) \quad(4.31-4.69) \end{gathered}$ | 0.174 | $1.71{ }^{\text {ns }}$ | -2.52* |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Eleven Point | M | 11 | $\begin{gathered} \mathrm{GR}=-0.0200 \mathrm{TL}+7.3704 \\ (-.0012--.0395) \quad(6.80-7.94) \end{gathered}$ | 0.464 | $-2.06{ }^{\text {ns }}$ | -2.42* |
|  | Both | 46 | $\begin{gathered} \mathrm{GR}=-0.0127 \mathrm{TL}+5.2193 \\ (-.0072--.0184) \quad(4.96-5.48) \end{gathered}$ | 0.354 | $0.428{ }^{\text {ns }}$ | -4.54*** |
| Gasconade | M | 31 | $\begin{array}{r} \mathrm{GR}=-11.2865 \mathrm{LogTL}+30.2577 \\ (-5.78--16.27) \quad(30.06-30.46) \end{array}$ | 0.561 | $1.76{ }^{\text {ns }}$ | -4.96*** |
|  | F | 53 | $\begin{gathered} \mathrm{GR}=-0.0085 \mathrm{TL}+4.2992 \\ (-.0051--.0121) \quad(4.20-4.39) \end{gathered}$ | 0.346 | $1.60{ }^{\text {ns }}$ | -4.87*** |
| Big Piney | M | 19 | $\begin{gathered} \mathrm{GR}=-0.0101 \mathrm{TL}+4.4964 \\ (-.0038--.0159) \quad(4.31-4.69) \end{gathered}$ | 0.512 | $1.65{ }^{\text {ns }}$ | -3.20** |
|  | F | 25 | $\begin{gathered} \mathrm{GR}=-0.0122 \mathrm{TL}+5.7711 \\ (-.0035--.0207) \quad(5.47-6.07) \end{gathered}$ | 0.311 | $2.06{ }^{\text {ns }}$ | -2.84** |

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To compare the regression coefficient estimates of the TLgrowth relationships (Table 16), each growth was divided by the standard deviation of the growths for a particular regression to equalize variances (Murphy, 1973). The Big Piney and Gasconade growth equations used in the comparisons are those of the females. Data for the growth equations of the North Fork and Niangua hellbenders are from Peterson et al. (1983) and Taber et al. (1975), respectively. The regressions for the Big Piney, Gasconade, and Spring Rivers are not influenced by juvenile growth because no juveniles were recaptured or because there was differential growth of the sexes. Thus they are not as likely to be as steep as the growth lines of the other three populations which are influenced by juvenile growth (Fig. 4). Because slopes are significantly different in only one test, but elevations are different in all tests, growth per size among adults is greatest in the Niangua population followed by the Spring River, Gasconade, North Fork, Big Piney, and Eleven Point hellbenders. All possible comparisons were not made for reasons given previously with regard to length-fecundity comparisons. Length-weight relationships. A power function was fitted to the length-weight data of recaptured adults of both sexes (Table 17). The Big Piney females exhibit a significant nonlinear component; however, the measure of the amount of variation in weight explained by the regression on $T L$, the coefficient of variation, is large. Between sexes there are no significant differences in the slopes of the length-weight relationships but significant differences in the $y$-intercepts. This suggests that females are larger than males in the Eleven Point and Spring Rivers as in the Big Piney and
Table 16. T-tests of the significance of the difference between slopes and $y$-intercepts of the regressions of growth on TL as described by Simpson et al. (1960). Data for the Niangua and North Fork regressions are from Taber et al. (1975) and Peterson et al. (1983). Growths were transformed by division of the growths by their standard deviation to equalize variances. The Gasconade and Big Piney regressions are those of the females only.

| Test | N | Variances $\qquad$ F | $\frac{\text { Slopes }}{t}$ | $\frac{\text { Elevations }}{\mathrm{t}}$ | Conclusions |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | variances | slopes | elevations |
| Spring vs. Niangua | 44;260 | $1.297^{\text {ns }}$ | $1.184^{\text {ns }}$ | -13.076*** | equal | equal | different |
| Spring vs. Gasconade | 44;53 | $1.272^{\text {ns }}$ | $0.619^{\text {ns }}$ | -27.955*** | equal | equal | different |
| Gasconade vs. North Fork | 53;99 | $1.384^{\mathrm{ns}}$ | 3.419*** | - | equal | different | - |
| Niangua vs. North Fork | 260;99 | $1.357^{\text {ns }}$ | $0.857^{\text {ns }}$ | 9.706*** | equal | equal | different |
| Big Piney vs. Gasconade | 25;53 | $1.018^{\mathrm{ns}}$ | $0.285^{\text {ns }}$ | -18.730*** | equal | equal | different |
| Big Piney vs. North Fork | 25;99 | $1.389^{\text {ns }}$ | $1.010^{\mathrm{ns}}$ | -15.767*** | equal | equal | different |
| Eleven Point vs. <br> Big Piney | 46;25 | $1.019^{\mathrm{ns}}$ | $0.083{ }^{\text {ns }}$ | 7.684*** | equal | equal | different |

[^9]```
Figure 4. A comparison of the length-growth regressions of
    hellbenders: Spring River (S), Eleven Point (EP),
    Gasconade females, (G), Big Piney females (BP),
    Niangua (N), North Fork (NF). Lines for the
    Niangua and North Fork hellbenders are from
    Peterson et al. (1983). C. a. bishopi are
    represented by dashed lines, c. a. alleganiensis
    by solid lines.
```

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Table 17. Regression of the logarithm of weight (WT) on the logarithm of total length (TL) of recaptured hellbenders using Bartlett's (1949) method as described by Simpson et al. (1960) and a comparison of the regression coefficients between sexes of each population.

| Population | Sex | N | Regression equation | $r^{2}$ | t value |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | nonlinear component | slope differing from zero | slopes | elevations |
| Spring | M | 38 | $W T=0.0000532 \mathrm{TL}{ }^{2.6998}$ | 0.756 | $-1.834^{\text {ns }}$ | 5.151*** | $-0.022^{\mathrm{ns}}$ | 34.024*** |
|  | F | 36 | $W \mathrm{~T}=0.0000005 \mathrm{TL}^{3.4675}$ | 0.716 | $-1.530^{\mathrm{ns}}$ | 7.041*** |  |  |
| Eleven Point | M | 12 | $\begin{aligned} & W T=0.0000282 \mathrm{TL}^{2.7764} \\ & W T=0.0000160 \mathrm{TL}^{2.8770} \end{aligned}$ | 0.931 | $-0.185^{\text {ns }}$ | 4.352** | $-0.005^{\text {ns }}$ | 2.355* |
|  | F | 27 |  | 0.911 | $-0.455^{\text {ns }}$ | 4.264*** |  |  |
| Gasconade | M | 39 | $W \mathrm{~T}=0.0000492 \mathrm{TL}{ }^{2.7003}$ | 0.905 | $-0.013^{\text {ns }}$ | 7.777*** | $-0.049^{\text {ns }}$ | 11.819*** |
|  | F | 68 | WT $=0.0000129 \mathrm{TL}^{2.9048}$ | 0.800 | $0.086{ }^{\text {ns }}$ | 10.228*** |  |  |
| Big Piney | MF | $\begin{aligned} & 31 \\ & 50 \end{aligned}$ | $\begin{aligned} & W T=0.0000436 \mathrm{TL}^{2.7027} \\ & W T=0.0000011 \mathrm{TL}^{3.3121} \end{aligned}$ | $\begin{aligned} & 0.838 \\ & 0.877 \end{aligned}$ | $0.677^{\text {ns }}$ | 7.298*** | $-0.035^{\mathrm{ns}}$ | 30.388*** |
|  |  |  |  |  | -2.196* | 11.179*** |  |  |

*, **, *** $\mathrm{P}<0.05,0.01,0.001$ respectively

Gasconade Rivers.
Maximum size. The average maximum TL or asymptotic length predicted by each length-growth regression is the $x$-intercept. Unlike the Spring and Eleven Point populations, a significant difference in growth of sexes was detected in the Gasconade and Big Piney populations. Thus asymptotic length can be predicted for each sex in the latter two rivers (Table 18). The asymptotic length was substituted into the length-weight relationships to extimate an average maximum weight.

Environmental Factors Influencing Growth

Temperature. The heavily spring-influenced streams, the Spring and Eleven Point Rivers, remained fairly stenothermal and exhibited similar temperatures (Figs. 5, 6, and 7). However, growth of hellbenders was about the most disparate in these two rivers. The North Fork, also highly influenced by springs, showed less variation in temparatures than the Big Piney and Gasconade, remaining warmer in winter and cooler in summer like the other south-flowing streams. The Niangua River exhibited rather wide variation despite temperature recordings being made only about 4 km downstream from Bennett Springs, the fourth largest spring in Missouri (Vineyard and Feder, 1974).

Relative crayfish abundance. Catch per unit effort of crayfish was determined when conditions were "good" (Table 19) as described for catch per unit effort of hellbenders. Crayfish of all sizes and species (Orconectes spp.) were captured. The North Fork hellbender biomass is from Peterson et al. (1983). No information on hellbender biomass is available from the site of collection
Table 18. Predicted average maximum total length (TL) from lengthgrowth regressions in Table 15 and the expected weight (WT) from length-weight regressions in heaviest individuals captured.

| Population | Average predicted maximum size |  | Longest observed individual |  | Heaviest observed individual |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \mathrm{TL} \\ (\mathrm{~mm}) \end{gathered}$ | $\begin{aligned} & \text { WT } \\ & (\mathrm{g}) \end{aligned}$ | $\begin{gathered} \mathrm{TL} \\ (\mathrm{~mm}) \end{gathered}$ | $\begin{array}{r} \text { WT } \\ (\mathrm{g}) \end{array}$ | $\begin{gathered} \mathrm{TL} \\ (\mathrm{~mm}) \end{gathered}$ | $\begin{aligned} & \text { WT } \\ & (\mathrm{g}) \end{aligned}$ |
| Spring |  |  |  |  |  |  |
| Male | $559{ }^{\text {a }}$ | 1392 | 600 | 1656 | 600 | ${ }^{1656}+$ |
| Female |  | 1593 | 596 | 1872 | 578 |  |
| Eleven |  |  |  |  |  |  |
| Point |  |  |  |  |  |  |
| Male | $411{ }^{\text {a }}$ | 510 | 444 | 565 | 444 | 565 |
| Female |  | 530 | 446 | 784 | 446 | 784 |
| Gasconade |  |  |  |  |  |  |
| Male | 480 | 723 | 515 | 857 | 487 | 858 |
| Female | 506 | 923 | 548 | 1154 | 543 | 1584 |
| Big Piney |  |  |  |  |  |  |
| Male | 445 | 627 | 502 | 953 | 502 | 953 |
| Female | 473 | 797 | 523 | 878 | 476 | 1018 |

${ }^{\text {No significant difference was detected in the growth of the sexes }}$

+ Only individual that weighed more than 2110 g , the maximum the triple-beam balance could measure

```
Figure 5. Maximum and minimum temperatures and current
temperature when thermometers were checked for
the Spring (.), Eleven Point (O), Gasconade (*),
and Big Piney (t) Rivers from fall 1980 to fall
1981.
```

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```
Figure 6. Maximum and minimum temperatures for the cold
    season months, November through March 1980-81.
    Vertical lines are the range, crossbars the
    mean, and boxes represent }\pm1\mathrm{ standard deviation.
    C. a. bishopi rivers: Spring (S), Eleven Point
    (EP), North Fork (NF). C. a. alleganiensis
    rivers: Gasconade (G), Big Piney (BP),
    Niangua (N).
```

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Figure 7. Maximum and minimum temperatures for the warm
season months, April through October, 1981.

Vertical lines are the range, crossbars the
mean, and boxes represent $\pm 1$ standard deviation.
C. a. bishopi rivers: Spring (S), Eleven Point
(EP), North Fork (NF). C. a. alleganiensis
rivers: Gasconade (G), Big Piney (BP),

Niangua (N).

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Table 19. Catch per unit effort of crayfish (Orconectes spp.)
made only during summer. Hellbender biomass is from Table
9, except the North Fork datum is from Peterson et al. (1983).

|  | Average catch <br> per hour |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Population | Site | weight(g) no. | Man-hours | Hellbender <br> biomass |  |
| Spring | 1 | 1792 | 177 | 1.0 | 396 |
|  | 2 | 888 | 127 | 1.0 | 74 |
| Eleven | 1 | 1100 | 102 | 0.9 | 92 |
| Point | 2 | 729 | 92 | 1.0 | 209 |
| North Fork | 1 | 1510 | 180 | 1.0 | 157 |
| Gasconade | 1 | 349 | 50 | 1.3 | 59 |
| Big Piney | 1 | 1165 | 88 | 1.0 | 112 |
|  | 2 | 923 | 153 | 1.1 | 334 |
| Niangua | - | 2721 | 182 | 1.0 | - |

of crayfish in the Niangua; however, catch per time indicates a dense population: an average of 17.2 hellbenders were captured per hour based on 1.9 hours of effort, a greater catch than in any other river. By wet wieght the crayfish catch correlation with hellbender biomass is not significant ( $r=0.534, \mathrm{P}>0.1$ ), but the number of crayfish caught per hour is weakly correlated with standing crop of hellbenders $(r=0.657,0.1>P>0.05)$. If the average estimate of hellbender biomass made by Topping (personal communication) for 10 Niangua collection sites ( $871 \mathrm{~kg} / \mathrm{ha}$ ) is divided in half as an approximate measure of biomass for the entire area of the sites (not just diurnal habitat) and paired with the Niangua crayfish catch in Table 19, the result is a correlation coefficient significantly different from zero for weight of crayfish caught per hour versus hellbender biomass ( $\mathrm{r}=0.730, \mathrm{P}<0.05$ ) as well as number of crayfish caught per hour versus hellbender biomass $(r=0.731, P<0.05)$.

To determine the error that could be involved in the crayfish catch, three other collectors besides the author worked site 1 in the Eleven Point concurrently for approximately one hour. For the four efforts the mean and standard error of the number of crayfish caught per hour was 114.4 and 11.2 crayfish, respectively. The mean and standard error of the crayfish weight caught per hour was 1119.2 and 142.1 g , respectively. This suggests that catch per unit effort may be a useful method for assessing relative crayfish abundance.

Age
The length-growth regressions can be integrated to generate an equation relating age in months to TL in mm . If age at any specific size is known, the constant of integration can be calculated. For the Eleven Point hellbenders, age and size at metamorphosis, 18 months and 125 mm TL according to Goin and Goin (1971), was used to calculate the constant. This reference point was also used by Peterson et al. (1983) and Taber et al. (1975) for their growth models. The length-age curve in Fig. 8 was then generated from this length-age equation: Age ${ }_{(m o n t h s)}=(1 /-0.0127) \mathrm{ln}$ $(5.2193+(-0.0127) T L(\mathrm{~mm}))+119.5534$. Reaching sexual maturity at about 300 mm TL, Eleven Point females are approximately seven to eight years old.

Because juvenile growth was not measured in the other three populations, transformation size and age was considered too far beyond the scope of the growth regression to be useful. Therefore size and age at sexual maturity for females was estimated to calculate the integration constant for the Spring, Gasconade, and Big Piney populations. Because maturity for females is attained when about seven to eight years old in the Eleven Point and also in the North Fork and Niangua populations (Peterson et al., 1983; Taber et al., 1975), this age is assumed for the other three populations. Size at maturity is about 370 mm TL for Big Piney females. This size is also assumed for the Gasconade and Spring River females because their growth is slightly less than that of the Niangua females which mature when about 38 to 39 cm TL (Ingersol, 1982; Taber et al., 1975). Thus 84 months and 370 mm TL were used to

```
Figure 8. Length-age curve for Eleven Point hellbenders.
    The line is the integral form of the regression
    in Fig. }3\mathrm{ with the integration constant calculated
    by substitution of the estimated size (125 mm TL)
    and age (18 months) at metamorphosis, according
    to Goin and Goin (1971). Division by 12 changes
    age from months to years.
```

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generate length-age curves (Fig. 9) from the following equations for the Spring, Gasconade, and Big Piney adult females, respectively:
 Age ${ }_{(\text {months })}=(1 /-0.0122) \ln (4.2992+(-0.0085) \mathrm{TL}(\mathrm{mm}))+100.8715$, Age $_{(\text {months })}=(1 /-0.0122) \ln \left(5.7711+(-0.0122) \mathrm{TL}_{(\mathrm{mm})}\right)+102.7547$. Age-specific Fecundity and Survivorship

Fecundity tables. Because the method of aging is crude, intervals of less than five years were not considered for aging the adult females sampled for age-specific fecundity information (Table 20). The number of females of the zero age class was estimated: by summing the products of columns two and four for each age interval, then dividing by two. This method assumes that all eggs are laid and fertilized and that the sex ratio at conception is $1: 1$.

Although fecundity increases with size and thus age as evident from the length-fecundity relationships (Fig. 2), the greatest number of eggs laid each year by the population is produced by the young adults. This is because fewer animals are alive in each subsequent five-year age interval. This is particularly so in the Eleven Point, less so in the Gasconade and Spring Rivers where the populations are dominated to a greater degree by older animals. The number of female zygotes produced per female varies among populations for females of similar age (column 6). An approximate Hartley's test indicated no significant difference in the variance of the female zygotes produced per female among populations $(H=2.522, r=4, n=5.25, P>0.05)$. A single factor ANOVA did indicate a significant difference in the means of the number of female zygotes

```
Figure 9. Length-age curve for adult females: Spring
    River (S), Eleven Point (EP), Gasconade (G),
    Big Piney (BP). The lines represent the
    integral form of the regressions in Fig. 3.
    Except for the Eleven Point, the integration
    constant was calculated by substitution of the
    estimated size ( }370\textrm{mm TL}\mathrm{ ) and age (84 months)
    at sexual maturity.
```

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Table 20. Number of female zygotes produced per female over age intervals for all
ages from sexual maturity to hypothesized oldest age.

| $\begin{aligned} & \text { Age in } \\ & \text { years } \\ & \text { (interval) } \end{aligned}$ | Spring River |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Sampled frequency of females late summer to fall 1980 | Expected TL in mm from age-length relationship | Expected no. of eggs produced per female from lengthfecundity relationship | ```Product of column 2 and column 4``` | No. of female zygotes per female, column 4 $\div 2$ |
| 0 | 29,170 ${ }^{\text {a }}$ | - | - | - | - |
| $\begin{gathered} 11.5 \\ (7-16) \end{gathered}$ | 39 | 435 | 471 | 18,369 | 235.5 |
| $\begin{gathered} 19 \\ (17-21) \end{gathered}$ | 22 | 498 | 619 | 13,618 | 309.5 |
| $\begin{gathered} 24 \\ (22-26) \end{gathered}$ | 13 | 521 | 673 | 8,749 | 336.5 |
| $\begin{gathered} 29 \\ (27-31) \end{gathered}$ | 6 | 535 | 705 | 4,230 | 352.5 |
| $\begin{gathered} 41.5 \\ (32-51) \end{gathered}$ | 18 | 551 | 743 | $\frac{13,374}{58,340}$ | $\begin{gathered} 371.5 \\ =29,170^{\mathrm{a}} \end{gathered}$ |

[^10]Table 20 cont. Number of female zygotes produced per female over age intervals for all ages from sexual maturity to hypothesized oldest age.

| Age in years (interval) | ```Sampled frequency late summer to fall 1980``` | Eleven <br> Expected TL in mm from age-length relationship | t River <br> Expected no. of eggs produced per female from lengthfecundity relationship | ```Product of column 2 and column 4``` | No. of female zygotes per female, column 4 $\div 2$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | 8,329 ${ }^{\text {a }}$ | - | - | - | - |
| $\stackrel{9}{(7-11)}$ | 38 | 320 | 206 | 7,828 | 103 |
| $\begin{gathered} 14 \\ (12-16) \end{gathered}$ | 17 | 368 | 294 | 4,998 | 147 |
| $\begin{gathered} 19 \\ (17-21) \end{gathered}$ | 7 | 391 | 337 | 2,359 | 168.5 |
| $\begin{gathered} 24 \\ (22-26) \end{gathered}$ | 1 | 402 | 357 | 357 | 178.5 |
| $(27-51)$ | 3 | 410 | 372 | 1,116 | 186 |
|  |  |  |  | 16,658 * | 8,329 ${ }^{\text {a }}$ |

$a_{\text {Assumes }}$ a $1: 1$ zygote sex ratio and that all eggs are laid and fertilized
Table 20 cont. Number of female zygotes produced per female over age intervals for
all ages from sexual maturity to hypothesized oldest age.

|  |  | Gasconade River |
| :---: | :---: | :---: | :---: | :---: | :---: |

$a_{\text {Assumes }}$ a $1: 1$ zygote sex ratio and that all eggs are laid and fertilized
Table 20 cont. Number of female zygotes produced per female over age intervals for
all ages from sexual maturity to hypothesized oldest age.

|  |  | Big Piney River |
| :--- | :---: | :---: | :---: | :---: | :---: |

[^11]produced per female among the four populations ( $F=13.41, \mathrm{P}<0.001$ ). The test is not totally conclusive because of the slightly different age classes in each population.

Even though for a particular size the Eleven Point and Spring River females produce similar numbers of eggs (Table 13), the agespecific production is much greater by the Spring River females. This is because their growth is greater than that of the Eleven Point animals. Likewise, the age-specific production of eggs by Gasconade females is greater than that of Big Piney females.

Life tables. Survivorship was determined from the sample of females utilized in the fecundity tables and aged based on TL (Table 21). Therefore a stable age distribution was assumed for each population, i. e., survivorship and fecundity rates were assumed to have been constant, as well as rate of increase (assumed equal to zero in this case), for at least a few generations. Survivorship follows the pattern of Deevey's (1947) Type III survivorship curve with high mortality of the young but relatively high survival rates for adults in all populations (Fig. 10). An approximate Hartley's test indicated no significant difference in variances of the survivorship rates (column 6 figures) among the four populations $(\mathrm{H}=2.18, \mathrm{r}=4, \mathrm{n}=5.25, \mathrm{P}>0.05)$. ANOVA indicated no significant difference in the means of the survivorship rates ( $\mathrm{F}=0.701, \mathrm{P}>0.05$ ).

Female Cohort Biomass and Production
From 74,363 eggs, half the annual fecundity of the estimated 250 adult Spring River females in the two collection sites, about 186 kg of salamander biomass are produced over about 51 years in the Spring River (by one cohort), or approximately $15 \mathrm{~g} / \mathrm{yr}$ wet
Table 21. Age-specific survivorship of females. Replacement rate is assumed to equal 1.0 .

| Age in years (interval) | Spring River |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Sampled frequency of females late summer to fall 1980 |  | ```Proportion dying during age X to X+1``` | $\begin{aligned} & \text { Mortality } \\ & \text { rate } \end{aligned}$ | $\begin{aligned} & \text { Survivorship } \\ & \text { rate } \end{aligned}$ |
| 0 | 29, $170^{\text {a }}$ | 1.0000000 | 0.9986630 | 0.9986630 | 0.0013370 |
| $\begin{gathered} 11.5 \\ (7-16) \end{gathered}$ | 39 | 0.0013370 | 0.0005850 | 0.4375467 | 0.5624533 |
| $\begin{gathered} 19 \\ (17-21) \end{gathered}$ | 22 | 0.0007520 | 0.0003063 | 0.4073138 | 0.5926862 |
| $\begin{gathered} 24 \\ (22-26) \end{gathered}$ | 13 | 0.0004457 | 0.0002400 | 0.5384788 | 0.4615212 |
| $\begin{gathered} 29 \\ (27-31) \end{gathered}$ | 6 | 0.0002057 | 0.0000514 | 0.2498788 | 0.7501215 |
| $\begin{gathered} 41.5 \\ (32-51) \end{gathered}$ | 18 | 0.0001543 | - |  | - |
| Mean Generation Time $=23.3$ years |  |  |  |  |  |

[^12]Table 21 cont. Age-specific survivorship of females. Replacement rate is assumed to
equal 1.0 .

| Age in years (interval) | Sampledfrequencyof femaleslate summerto fall1980 | Eleven <br> No. surviving as a fraction of zygotes | River | $\begin{aligned} & \text { Mortality } \\ & \text { rate } \end{aligned}$ | Survivorship rate |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | ```Proportion dying during age X to X+1``` |  |  |
| 0 | $8,329{ }^{\text {a }}$ | 1.0000000 | 0.9954376 | 0.9954376 | 0.0034320 |
| $\stackrel{9}{(7-11)}$ | 38 | 0.0045624 | 0.0025213 | 0.5526258 | 0.4473742 |
| $\begin{gathered} 14 \\ (12-16) \end{gathered}$ | 17 | 0.0020411 | 0.0012007 | 0.5882612 | 0.4117388 |
| $\begin{gathered} 19 \\ (17-21) \end{gathered}$ | 7 | 0.0008404 | 0.0007203 | 0.8570919 | 0.1429081 |
| $\begin{gathered} 24 \\ (22-26) \end{gathered}$ | 1 | 0.0001201 | 0.0000481 | 0.4004996 | 0.5995004 |
| $\begin{gathered} 39 \\ (27-51) \end{gathered}$ | 3 | 0.0000720 | - | - | - |
| Mean Generation time $=14.2$ years |  |  |  |  |  |

$\mathrm{a}_{\text {From }}$ fecundity table
Table 21 cont. Age-specific survivorship of females. Replacement rate is assumed to
equal 1.0 .

| Age inyears(interval) | Sampledfrequencyof femaleslate summerto fall1980 | Gascona <br> No. surviving as a fraction. of zygotes | River | $\begin{aligned} & \text { Mortality } \\ & \text { rate } \end{aligned}$ | Survivorship rate |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | ```Proportion dying during age X to X+1``` |  |  |
| 0 | 17,975 ${ }^{\text {a }}$ | 1.0000000 | 0.9984423 | 0.9984423 | 0.0015577 |
| $\stackrel{9}{(7-11)}$ | 28 | 0.0015577 | 0.0005007 | 0.3214354 | 0.6785646 |
| $\begin{gathered} 14 \\ (12-16) \end{gathered}$ | 19 | 0.0010570 | 0.0002781 | 0.2631031 | 0.7368969 |
| $\begin{gathered} 19 \\ (17-21) \end{gathered}$ | 14 | 0.0007789 | 0.0002226 | 0.2857876 | 0.7142124 |
| $\begin{gathered} 24 \\ (22-26) \end{gathered}$ | 10 | 0.0005563 | 0.0001669 | 0.3000180 | 0.6999820 |
| $\begin{gathered} 29 \\ (27-31) \end{gathered}$ | 7 | 0.0003894 | 0.0001391 | 0.3572162 | 0.6427838 |
| $\begin{gathered} 41.5 \\ (32-51) \end{gathered}$ | 18 | 0.0002503 | - - | $\text { ime }=23.4$ | ears |

$\mathrm{a}_{\text {From }}$ fecundity table

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Table 21 cont. Age-specific survivorship of females. Replacement rate is assumed to
equal 1.0 .

| Age in years (interval) | Big Piney River |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | ```Samples frequency of females late summer to fall 1980``` |  | Proportion <br> dying <br> during age <br> X to $\mathrm{X}+1$ | $\begin{aligned} & \text { Mortality } \\ & \text { rate } \end{aligned}$ | Survivorship rate |
| 0 | 15,024 ${ }^{\text {a }}$ | 1.0000000 | 0.9975373 | 0.9975373 | 0.0024627 |
| $\begin{gathered} 9 \\ (7-11) \end{gathered}$ | 37 | 0.0024627 | 0.0006656 | 0.2702725 | 0.7297275 |
| $\begin{gathered} 14 \\ (12-16) \end{gathered}$ | 27 | 0.0017971 | 0.0008653 | 0.4814980 | 0.5185020 |
| $\begin{gathered} 19 \\ (17-21) \end{gathered}$ | 14 | 0.0009318 | 0.0007321 | 0.7856836 | 0.2143164 |
| $\begin{gathered} 24 \\ (22-26) \end{gathered}$ | 3 | 0.0001997 | 0.0000266 | 0.1331998 | 0.8668002 |
| $\begin{gathered} 39 \\ (27-51) \end{gathered}$ | 13 | 0.0001731 | an Generat | $\text { ime }=18.6$ | - |

[^13]```
Fibure 10. Age-specific survivorship of a cohort of 10,000
    females based on Table 21 and age-specific
    fecundity from Table 20: Spring River (S),
    Eleven Point (EP), Gasconade (G), Big Piney
    (BP).
```

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weight from each clutch of female zygotes is consumed by predators and decomposers (Table 22). Similarly, for the Eleven Point, it is about $7 \mathrm{~g} / \mathrm{yr}$ per clutch of females. For the Gasconade and Big Piney populations, 7 g and 9 g of female biomass per year per clutch, respectively, are consumed. This production is highly influenced by the great early mortality as evident from the first value in the Ix column in all populations. The relatively large biomass value for the Spring River is due to the greater growth rate (Fig. 4) and particularly the large fecundity of the Spring River females (Figs. 2 and 10).

```
If a species produces a new generation each year so that overlapping cohorts are present and the population has a stable age distribution, then annual production for any one year would be equal to one cohort production (Waters, 1977). Thus the sum of the weights of individuals dying between all age intervals in Table 22 represents the annual production of female hellbenders ( \(P\) ) calculated by the summation-removal method in the area of the collection sites (Table 23). Yearly production in all collection sites was also calculated based on an estimate of instantaneous growth ( \(\mathrm{P}^{\prime}\) ).
As might be expected because of the influence of biomass in their calculation, the correlation between production estimates ( \(\mathrm{P}^{\prime}\) ) and number of crayfish caught per hour in each site (Fig. 11) is significant (r=0.821, \(\mathrm{P}<0.05\) ), but no significant correlation exists with weight of crayfish caught per hour ( \(r=0.647, \mathrm{P}>0.1\) ). Again, if production for site 1 in the North Fork (calculated by the instantaneous growth method) and half of Topping's (1981 and personal communication) average estimate of production by Niangua hellbenders
```

Table 22. Biomass furnished to the system's predators and decomposers from a cohort of Table 22. The number of adult females alive is the combined estimate of female population size for the two collection sites based on Tables 9 and 10 .

|  | No. alive Nx | Spring River |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Expected weight from length- weight relationship (grams) Wx | Weight increment between intervals (grams) Gx | Average individual growth per year (grams) | No. dying between intervals Dx | Average weight at death (grams) Ix | Total weight of individuals dying between intervals (grams) Tx |
| 0 | 74,363 | $0.1{ }^{\text {a }}$ | - | - | - | - | - |
| 7-16 | 100 | 668 | 668 | 58 | 74,263 | 1.0 | 74,263 |
| 17-21 | 56 | 1,067 | 399 | 53 | 44 | 891.4 | 39,222 |
| 22-26 | 33 | 1,248 | 181 | 36 | 23 | 1,173.7 | 26,995 |
| 27-31 | 15 | 1,369 | 121 | 24 | 18 | 1,303.0 | 23,454 |
| 32-51 | 46 | 1,516 | 147 | 12 | 15 | $1,481.7$ | 22,226 |
|  |  |  |  |  | 74,363 |  | 186,159 |

[^14]Table 22 cont. Biomass furnished to the system's predators and decomposers from a cohort of
females. The number of adult females alive is the combined estimate of female population size for the two collection sites based on Tables 9 and 10 .

| ```Age interval (years) x``` | No. alive Nx | Eleven Point River |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Expected weight from lengthweight relationship (grams) Wx | Weight increment between intervals (grams) Gx | Average individual growth per year (grams) | No. dying between intervals Dx | Average weight at death (grams) Ix | Total weight of individuals dying between intervals (grams) Tx |
| 0 | 14,626 | $0.1{ }^{\text {a }}$ | - | - | - | - | - |
| 7-11 | 59 | 258 | 258 | 29 | 14,567 | 1.1 | 16,024 |
| 12-16 | 26 | 386 | 128 | 26 | 33 | 314.4 | 10,375 |
| 17-21 | 11 | 459 | 73 | 15 | 15 | 416.9 | 6,254 |
| 22-26 | 2 | 498 | : 39 | 8 | 9 | 466.1 | 4,195 |
| 27-51 | 4 | 527 | 29 | 2 | 2 | 509.6 | 1,019 |
|  |  |  |  |  | 14,626 |  | 37,867 |
| ${ }^{\text {Approx }}$ <br> IX $=W x$ | ximate $x-1+$ | ight of an egg $x / N x-1) \cdot G x) ;$ | efore layin $x=D x \cdot I$ | from Spight | 1967) |  |  |

Table 22 cont. Biomass furnished to the system's predators and decomposers from a cohort of females. The number of adult females alive is the estimate of female population size for the collection site based on Tables 9 and 10 .

| Age interval (years) x | No. alive Nx | ```Expected weight from length- weight relationship (grams) Wx``` | Gasconade River |  | No. dying between intervals Dx | Average weight at death (grams) Ix | Total weight of individuals dying between intervals (grams) TX |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Weight increment between intervals (grams) Gx | Average individual growth per year (grams) |  |  |  |
| 0 | 25,110 | $0.1{ }^{\text {a }}$ | - | - | - | - | - |
| 7-11 | 39 | 450 | 450 | 50 | 25,071 | 0.8 | 20,057 |
| 12-16 | 26 | 611 | 161 | 32 | 13 | 557.3 | 7,245 |
| 17-21 | 20 | 727 | 116 | 23 | 6 | 700.2 | 4,201 |
| 22-26 | 14 | 802 | 75 | 15 | 6. | 779.5 | 4,677 |
| 27-31 | 10 | 846 | 44 | 9 | 4 | 833.4 | 3,334 |
| 32-51 | 25 | 902 | 56 | 4 | 10 | 881.0 | 8,810 |
|  |  |  |  |  | 25,110 |  | 48,324 |

[^15]Table 22 cont. Biomass furnished to the system's predators and decomposers from a cohort of females. The number of adult females alive is the combined estimate of female population size for the two colleciton sites based on Tables 9 and 10.

| Age interval (years) x | No. alive Nx | ```Expected weight from length- weight relationship (grams) Wx``` | Big Piney River |  | No. dying between intervals Dx | Average weight at death (grams) Ix | Total weight of individuals dying between intervals (grams) Tx |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Weight increment between intervals (grams) Gx | Average individual growth per year (grams) |  |  |  |
| 0 | 21,702 | $0.1{ }^{\text {a }}$ | - | - | - | - | - |
| 7-11 | 54 | 443 | 443 | 49 | 21,648 | 1.2 | 25,978 |
| 12-16 | 39 | 609 | 166 | 33 | 15 | 562.9 | 8,444 |
| 17-21 | 20 | 702 | 93 | 19 | 19 | 656.7 | 12,477 |
| 22-26 | 4 | 749 | 47 | 9 | 16 | 711.4 | 11,382 |
| 27-51 | 19 | 792 | 43 | 3 | 4 | 789.8 | 3,159 |
|  |  |  |  |  | 21,702 |  | 61,439 |

[^16]Table 23. Production of females ( $P$ ) by the summation-removal method assuming a stable population and production of both sexes ( $\mathrm{P}^{\prime}$ ) by the
instantaneous growth method, both are in $\mathrm{kg} / \mathrm{ha/} \mathrm{yr}$. The denominators
of the production : biomass ratios are the estimated biomass of females
(B) in the population based on data in Tables 9 and 10 and estimated
biomass of all hellbenders ( $\mathrm{B}^{\prime}$ ) in the site from Table 9 .

| Population | Site | $P$ | $P / B$ | $P^{\prime}$ | $P^{\prime} / B^{\prime}$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Spring | 1 | 96.5 | 0.49 | 77.3 | 0.20 |
|  | 2 |  | 2.61 | 14.0 | 0.19 |
| Eleven | 1 |  | 1.59 | 15.9 | 0.18 |
| Point | 2 | 92.4 | 0.71 | 36.7 | 0.18 |
| Gasconade | 1 | 22.6 | 0.59 | 11.2 | 0.19 |
| Big Piney | 1 |  | 1.14 | 20.3 | 0.18 |
|  | 2 | 73.1 | 0.38 | 60.4 | 0.18 |

Figure 11. The relationship between hellbender production
( $P^{\prime}$ ) in Table 23 and average crayfish catch per
hour in Table 19 for each of the seven collection
sites. Crayfish number per hour (©). Crayfish
weight per hour (\$).

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are paired with crayfish catch in Table }19\mathrm{ for the two rivers,
both correlations are significant: weight of crayfish caught
versus hellbender production ( }\textrm{r}=0.770,\textrm{P}<0.02\mathrm{ ) and number of cray-
fish caught versus hellbender production (r=0.733, P<0.05).
```

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## DISCUSSION

Population Size
Equal catchability. According to Caughley (1977), markrecapture experiments are divided into two classes: (1) those experiments involved with movement, growth rate, age-specific fecundity, or age-specific mortality and (2) those involved with estimating population size, rate of increase, rate of loss and gain, or rate of harvesting. The two classes differ in that in the former the behavior and number of animals in the unmarked segment of the population is irrelevant to estimating parameters, whereas in the second class of investigations, equal catchability of all individuals, marked and unmarked, is crucial to the accuracy of the estimates (Caughley, 1977). The design of this study was made with the last three components of the first class of experiments in mind. As many salamanders as possible were captured, and this may be contrary to random sampling. A mark-recapture model robust to unequal catchability has been developed for open populations (Pollock, 1982), but it requires considerably more sampling periods than are available from this study. However, since a mark-recapture study is time-consuming and expensive, it seems reasonable to gain all possible benefits from it as long as results are viewed with proper caution and the assumptions are tested.

Although no population exhibited highly significant departures from random sampling (Table 3), the Gasconade results were borderline in interpretation, insufficient data were available from the Eleven Point for Leslie's test, the Spring River data were
less numerous than desirable, and the Big Piney hellbenders differed significantly in catchability with regard to sexes and age classes (Tables 4 and 5). Thus all estimates resulting from the JollySeber analyses must be viewed with caution.

Eberhardt (1969) listed three causes for unequal catchability: (1) a property inherent in the individual, such as behavior near a capturing device, (2) the result of learning, becoming captureprone or capture-shy, and (3) a property depending on the relative opportunity of capture. As stated by Tilley (1980), it seems unlikely that salamanders will exhibit unequal catchability because of learning, but Beukema (1970) showed that the catchability of pike decreased considerably after being hooked previously, at least when using artificial bait. Some hellbenders, at least at certain times, do swim away quickly when their cover is removed making capture difficult. Whether this is an inherent property of some individuals or just expressed by all or most hellbenders only at certain times is unknown. The relative opportunity of capture does present a problem. Although hellbenders move about within a riffle (Coatney, 1982; Hillis and Bellis, 1971), there is some evidence that individuals may spend much of their time under the same cover during daylight hours (Coatney, 1982). Thus a hellbender that spends much time under an easily accessible, easily turned rock would more likely be captured by the handgrabbing method than a hellbender that spends much time in a deep pool under slab rock or a boulder too large to be moved. In the Big Piney, site 1 , several old large females were captured numerous times, one individual in nine of the ten samples, and resulted
in the large chi-square values in Tables 4 and 5. It could be that these large females that probably could defend their cover from other hellbenders spent much diurnal time under large flat, turnable rocks which were accessible and never bypassed, a rock that seemed very likely to shelter a hellbender. Therefore there exist some difficulties in the design of the investigation for population size estimation, and a few of the assumptions made by the Jolly-Seber method are questionable. However, according to Begon (1979), no actual mark-recapture study ever satisfies all assumptions in the very finest detail. Rather than rejecting all mark-recapture models because of this and deny any information, it is better to take a utilitarian view that the assumptions are valid so long as they do not lead to results which are so unreliable as to be useless. Begon (1979) stated that if the equal catchability assumption is violated the population size will be underestimated. Survivorship will also be underestimated (Carothers, 1973). But even with a statistically significant indication of unequal catchability, the bias in survivorship may be lower than that due to sampling variation and thus of little practical significance (Carothers, 1979). Further, Carothers (1979) stated that this points out the difficulty in relying only on significance tests to determine the appropriateness of mark-recapture models to a particular data set. Thus the estimates in Table 6 may at least be a minimum estimate of the numbers of hellbenders, and the effect on survivorship by possible nonrandom sampling may be negligible, despite Caughley's (1977) assertion that unequal catchability is the greatest source of error in mark-recapture studies.

```
There is significant evidence that Big Piney subgroups were not sampled randomly (Tables 4 and 5). Thus, because there were multiple recaptures of individuals in site 1 , frequency of capture models were compared to the observed frequency of recaptures to estimate the number of hellbenders in the site. Although Caughley (1977) refers to these methods as crude and insensitive, he stated that at least one of the three is liable to be a good fit. However, Wilbur and Landwehr (1974) believed that the frequency of capture models have very limited value based on their work with painted turtles.
The estimate of 806 hellbenders by the geometric distribution frequency of capture model seems too large, nearly 12 hellbenders per \(100 \mathrm{~m}^{2}\), especially in comparison to the catch per unit effort. This would indicate that only about 6 to \(10 \%\) of the population was captured during a sample, which does not seem reasonable. The frequency models assume no losses which is somewhat unrealistic just considering mortality.
Losses and gains to the population. The additions to the population (Table 6) probably do not reflect natality, as so few juveniles were captured. More likely the variation in the population estimates is due to sampling error as indicated by the standard errors, than to mortality or movement, although these are probably involved. Regression of the natural logarithm of the population size on time indicates little increase or decrease in all sites but one during the study. That one, site 1 in the Eleven Point, probably did not undergo a decrease in numbers of hellbenders, as no reason is known why such a decrease should occur. The slope
```

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of the regression was negative probably because of the few number
of estimates plus sampling error.
    The average annual survivorship rate for all collection
sites (0.27) determined from the pooled monthly survival rates
(MSRs) is very low compared to the survivorship indicated in the
life tables for adult females. This may be due partly to the
influence of juvenile survivorship on the MSRs. The difference
may be mostly due to sampling error and dispersal movements because
the survivorshiprates ( }\mp@subsup{P}{i}{}\mathrm{ ) are influenced by emigration as well
as mortality. The sites are not closed by natural boundaries,
and movements of several hundred meters upstream or downstream
by a few hellbenders have been reported (Wiggs, 1977; Nickerson
and Mays, 1973a). From Table 1 it is obvious that many hellbenders
captured once were never seen again, as is true of all hellbender
mark-recapture studies (Peterson et al., 1983; Taber et al., 1975;
Nickerson and Mays, 1973a; Hillis and Bellis, 1971). Further
as reported by Wiggs (1977), in the same site worked previously
many times, unbranded hellbenders are captured eachtime. This
suggests that there are many more hellbenders on the sites at
one time than the Jolly-Seber estimates would indicate, which
may be true if nonrandom sampling is a large problem, or it may
be that emigration and immigration are important factors in local
populations. Possibly both are involved. Thus the large estimate
by the frequency of capture method for site 1 in the Big Piney
may reflect a large hidden portion of the population and thus
    a poor sampling of each collection site, long-distance movements,
    or just the inexactness of the method.
```

Densities. Despite possible sources of error, such as differing current rates, depths, temperatures, visibilities, and, possibly, intensiveness of collecting in different streams, the relative abundance estimates, catch per unit effort, correlate well with absolute density estimates of hellbenders. This provides credibility for the Jolly-Seber estimates except in the Gasconade. In comparison to the other sites, the Gasconade site has very Jarge stretches of slab rock with generally wide, muddy shorelines where hellbenders are not found during the day, but these areas were considered in the absolute density estimates. Where there are large rocks, however, hellbenders are numerous. This is responsible for the large catch per unit effort, as shorelines were not searched and the smooth bedrock stretches were passed over quickly.

Nickerson and Mays (1973b) reported a density of 1 hellbender / $6-7 \mathrm{~m}^{2}$ to $1 / 13-16 \mathrm{~m}^{2}$ for one riffle in the North Fork River. However, only the area where hellbenders were found during daylight hours (the area where large rocks were present) was considered in the density estimates. In contrast, Peterson et al. (1983) estimated a density of 5.0 and 4.9 hellbenders / $100 \mathrm{~m}^{2}$ for two collection sites in the North Fork. Shoreline to shoreline area was considered in the estimates. Although there is merit in estimating only the area where hellbenders are located diurnally, slab rock and shoreline areas may be feeding areas at night (Beck, 1965). Also shoreline to shoreline area is less subjectively measured.

Hillis and Bellis (1971), based on a calculated mean activity radius of 10.5 m and a median activity radius of 6.0 m , estimated that the average circular home range of hellbenders in French Creek, Pennsylvania, was $346 \mathrm{~m}^{2}$ and the median home range was $113 \mathrm{~m}^{2}$. Coatney (1982) estimated a mean elliptical home range of $90 \mathrm{~m}^{2}$ for North Fork hellbenders. Because the estimates of density range from 1.0 to 6.3 hellbenders per $100 \mathrm{~m}^{2}$ in different sites (Table 9), home ranges likely overlap as reported by (Coatney, 1982). However, one hellbender followed by Coatney had a home range of just $18 \mathrm{~m}^{2}$ based on two weeks of telemetry data. It is uncommon to find more than one hellbender under a single rock (Nickerson and Mays, 1973a: Hillis and Bellis, 1971). However, Bishop (1941) reported finding three male hellbenders, a mass of hellbender eggs, and one Necturus under a single rock that measured just less than 1 m in width and length. While he was observing, a fourth male tried to enter the nest. Certainly if many large rocks are piled next to each other, several hellbenders sometimes can be uncovered in a few square meters during daylight hours. Coatney (1982) suggested that hellbenders whose home ranges overlap will not be in the area of overlap at night at the same time, at least in midsummer before the breeding season. Population Composition

Adults. External sexing of adults during late summer to fall is valid: 84 hellbenders were sexed externally as females and then dissected for fecundity and stomach contents; all vere female. Adult females, except in the Spring River, seem to be more numberous than adult males (Table 10). But the data for
the Big Piney is ambiguous because females are more catchable than males.

Reports of sex ratios for hellbenders vary from population to population and even from one time to another. Smith (1912) reported a male to female ratio of $2: 1$ to $3: 1$ based on several years of data. However, he found a higher proportion of females than males in "nonbreeding" areas during the summer and an equal number of the sexes during the breeding season. Hillis and Bellis (1917) reported a ratio of 1.58 males to 1 female. These two studies were made in Pennsylvania. In Missouri, Nickerson and Mays (1973a) reported male to female ratios based on midsummer collections for three rivers: North Fork, 1:1; Eleven Point, 1:2.45; Niangua, 1.3:1. Wiggs (1976) found a 1:1 ratio in the Niangua, but a report of 1 male to 1.6 females was made also for the Niangua and a ratio of 1:4 for hellbenders $15^{+}$years of age (Wiggs, 1977).

It seems reasonable that males, particularly older males are less numerous in some populations because adult male agespecific survivorship may be less than that of females (Peterson et al., 1983; Taber et al., 1975). Although differences in lengthspecific growth have not been detected between sexes in the North Fork (Peterson et al., 1983), Niangua (Taber et al., 1975), Spring, and Eleven Point Rivers, the growth of females is greater than that of males in both the Gasconade and Big Piney Rivers. Further, length-specific weight is greater for females in all the populations above, at least during midsummer to fall. It may be that the

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females' larger size may deter predation to a greater extent,
resulting in the greater number of females found in this study
and possibly greater survivorship. If such reasoning is correct,
why there exists an equal sex ratio in the Spring River is unknown.
The predators of adults are probably the same: large catfish
(Minton, 1972; Barbour, 1971), common water snakes, Nerodia sipedon
(Rhoads, 1895), turtles (Surface, 1913, as cited by Nickerson
and Mays, 1973a), other hellbenders (Wiggs, 1976; Nickerson and
Mays, 1973a), and perhaps mink, Mustela vison.
Juveniles. Other populations of hellbenders have been reported to be dominated by adults (Peterson et al., 1983; Taber et al., 1975; Hillis and Bellis, 1971), as in this study. Only a small number of larvae have been captured in other studies. Nickerson and Mays (1973a) captured about 2000 hellbenders in the North Fork River; eight of them were larvae. Taber et al. (1975) marked over 1100 Niangua hellbenders but only three larvae were captured. No larvae were captured by Peterson et al. (1983) who marked over 700 hellbenders in the North Fork. Undoubtedly the eggs and larvae and even the young subadults are subject to considerable predation. This agrees with the large young mortality in the life tables. Further, cannibalism of the eggs and young has been suggested to be an important factor in maintaining population stability (Mays and Nickerson, 1971).
Richmond (1965) stated that in those species that have great longevity and spend most of their lives as reproductive adults any stable or even slowly expanding population will consist almost entirely of mature animals. Further, he noted that such populations
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#### Abstract

are characterized by a very slow population turnover, an annual reproductive potential vastly greater than the annual replacement, and a high mortality of the young. Richmond (1965) argued that this is true of some turtles and suggested that this may be true of the hellbender. His speculation has been affirmed by this study and previous ones.


## Fecundity

Paired comparison of ovary oocyte number. No reason is known why the right ovary of Eleven Point and Gasconade females should contain more developing ova than the left ovary. Topping and Ingersol (1981) found that one ovary of both Niangua and North Fork females contained more oocytes than the other. They did not indicate whether the right or left ovary systematically contained a greater number of developing eggs.

Size at maturity. Female hellbenders appear to mature at different lengths in different populations. Eleven Point and Big Piney females mature at about 300 mm and 370 mm TL, respectively. Females seem to mature at a size between 330 and 380 mm TL in the North Fork (Peterson et al., 1983). Taber et al. (1975) believed Niangua females mature at about 380 mm . Ingersol (1982), who dissected both juvenile and adult females from the Niangua, found none less than 390 mm TL that was mature. The smallest mature female measured by Smith (1912) was 350 mm TL. Dundee and Dundee (1965) stated that C. a. bishopi females mature at about 238 mm snout-vent length or 330 mm TL based on females from three rivers and that C. a. alleganiensis females in Missouri attain maturity when about 247 mm snout-vent length. Males mature at a smaller

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size than females (Peterson et al., 1983; Taber et al., 1975;
Dundee and Dundee, 1965; Grenel1, 1939; Smith, 1912).
    TL-fecundity relationships. Despite a smiliar size-specific
production of eggs (Table 13), the actual fecundity of Spring
River females was much larger than that of Eleven Point females
(Fig. 10). This is because growth is much greater in the former
(Fig. 4), and fecundity increases with TL (Fig. 2). Similarly,
but to a lesser degree, Gasconade females produce more eggs than
Big Piney females. Comparisons of the regressions, including
those of Topping and Ingersol (1981), indicate that length-specific
fecundity is greater in C. a. bishopi than Missouri C. . a. alleganien-
sis (Table 13 and Fig. 2). Assuming, of course, that fecundity
is influenced by the genotype directly, this difference in fecundity
points to genetic differences between the subspecies, as all popu-
lations vary environmentally. Topping (personal communication)
had postulated that the greater North Fork female size-specific
fecundity over that of Niangua females might be an adaptation
that has been selected because of the smaller size of North Fork
hellbenders. Otherwise, the latter would have a much reduced
reproductive capacity in comparison to the larger Niangua females.
However, the greater growth of the Spring River C. a. bishopi
would seem a contradiction to the idea of selection for increased
fecundity in smaller C. a. bishopi over that of larger C. a.
alleganiensis. Yet if the environment does play an important
role in the growth measured in this study, the large growth of
Spring River hellbenders may be related to an enrichment of that
stream in the past few decades because of development by humans.
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Warren et al. (1964) found that the continuous enrichment of stream flow by a few mg / 1 of sucrose could increase the benthos and lead to a sevenfold or more increase in trout production. Thus the nutrient enrichment of the Spring River, as mentioned in the description of the study areas, might be an important factor in the growth and fecundity of Spring River hellbenders.

Spawning. According to Dundee and Dundee (1965), there is no complete separation of the spawning seasons of the two subspecies, but for the most part C. a. alleganiensis breed two months earlier. Eastern populations of $C$. a. alleganiensis do spawn earlier: from late August into September (Huheey and Stupka, 1967; Fitch, 1947; Bishop, 1941; King, 1939; Green, 1933; Smith, 1907), except in Alabama where breeding occurs in late September through early October (Mount, 1975). Spawning of C. a. alleganiensis in the Niangua River occurs from September into early October according to Ingersol (1982). However Dundee and Dundee (1965) reported collecting two ripe Niangua females on 14 November 1954 . On the same date they collected females in the Gasconade River that had ova 2 mm or less in diameter. It is possible that these females were spent, and the observed ooocytes were to be laid in subsequent seasons. Nickerson and Mays (1973a) observed egg laying by $\underline{C}$. a. bishopi in the North Fork on 13 September 1970 and 6 October 1973. Males were observed expelling milt on 3 November 1971, and one decomposing female was ripe. Dundee and Dundee (1965) reported that peak size of ova of Spring River C. a. bishopi is attained in late October, and Ratcliff (1965) reported that Spring River hellbenders breed in October. Baker (1963) stated that mature

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sperm were present in the testes of Spring River males in early
September; however, on 1 December the testes were considerably
enlarged and full of sperm. On }15\mathrm{ December no sperm were present,
and the gonads were reduced in size. It is thus unclear whether
the ripe condition of Spring River females collected on 12 December
in this study is unusual.
    Topping and Ingersol (1981) reported that Niangua females
may lay only 76% of their large ovarian eggs, and some females,
approximately 27.5% (Ingersol, 1982), do not lay any eggs in a
particular year but resorb all of them. They suggested that the
resorption of eggs might be a trait of other hellbender populations
because of the genetic uniformity described by Merkle et al. (1977).
Whether this occurs is unknown. The two ripe (in contrast to
1 3 \text { spent) Big Piney females killed 19 September may not have been}
going to spawn, but this date is early in the breeding season
for most hellbenders in the Ozarks. Smith (1912) also found hell-
benders with eggs that appeared to be degenerating. Swanson (1948)
killed five females in December, but none had more than five ova
left from the past spawning season.
                                    Stomach Contents
    Crayfish remains were the most frequently found food items
in stomachs in all previous studies throughout the range of Crypto-
branchus; fish remains were the second most frequently found items
(Wiggs, 1976; Nickerson and Mays, 1983a; Ferguson, 1961; Swanson,
1948; Bishop, 1941; Green, 1935; Netting, 1929; Smith, 1907; Reese,
1903). Wiggs, (1976) stated that because the digestion of crayfish
is slower (about one week, although chelae and chitinous cephalo-
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thoraxes may remain in the system for weeks) than of fish (two
to five days), the frequency of fish eaten in comparison to crayfish
may be two or three times greater than indicated by stomach contents.
Still, crayfish were found by Wiggs (1976) as the major stomach
item throughout the year with fish a distant second. She speculated
this was due to a large population of crayfish in the Niangua
River, the occupation of the same microhabitat as hellbenders by the
crayfish, and because crayfish were more easily captured by hellbenders
than were fish based on laboratory observations. Other food items
for adults are varied but apparently minor (cf. Nickerson and
Mays, 1973a). Possibly at certain times of the year, there are
exceptions, such as lampreys during their spawning (Nickerson
et al., 1983) or hellbender eggs as found in this study. Smith
(1912) reported that both sexes were voracious eaters of the eggs
and that the majority of adults collected during the spawning
season had stomachs filled with eggs. Conspecific oophagy has
been reported in a number of caudates, and it has been suggested
that the significance of such behavior should be studied in urodeles
(Kaplan and Sherman, 1980).
                                    Growth and Size Comparisons
    Growth models. Growth of individuals is variable so that
regressions explain only a relatively small amount of the variance
of mean monthly growth as evident from the coefficients of determina-
tion (Fig. 4). But similar regressions were developed with similar
variability about the regression lines by Peterson et al. (1983)
and Taber et al. (1975) for hellbenders in the North Fork and
Niangua Rivers, respectively. Two years after marking ended in
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#### Abstract

the North Fork and six years after branding ended in the Niangua, those populations were sampled, and long-term growth of recaptured individuals was compared to that predicted by the respective regressions. No significant difference in observed TLs and predicted TLs was found for North Fozk recaptures. Although final observed and predicted TLs differed significantly for Niangua recaptures, the growth model did predict growth without significant bias. This was true even though individuals were marked as much as 8 to 9 years previously (Peterson et al., 1985). Thus, except possibly the Spring River regression where the coefficient of determination is quite small, the regressions should be useful for predicting growth of individuals. Taber et al. (1975) and Peterson et al. (1983) discussed possible problems with the growth models. Differences in the growth of sexes was not detected by Peterson et al. (1983) or Taber et al. (1975) nor in the Spring and Eleven Point Rivers in this study. However, the Gasconade and Big Piney females exhibit greater length-specific growth than the males. Based on differences in the length-weight relationships in this study and the studies by Peterson et al. (1983) and Taber et al. (1975), it seems likely that female growth exceeds that of males in all populations in the Ozarks. But the greater elevation of the female length-weight regressions may at least in part reflect the growth of eggs. Because only adults can be sexed and only for a few months of the year, it is difficult to detect possible differences in the growth of sexes. Collins (1974) stated that females more often achieve a greater TL than males, but Bishop (1941) reported that males are often broader and heavier than


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females of the same length.
    Maximum size. Only the average maximum length is predicted
by each regression. The asymptotic length will underestimate
the TL of those individuals whose growth is greater than the general
trend described by the regression. This is true to a similar
degree for hellbenders in the North Fork and Niangua and the growth
models of Peterson et al. (1983) and Taber et al. (1975). Growth
of hellbenders is likely indeterminate, and the growth of very
large individuals is difficult to measure accurately because they
grow so slowly. Their growth like many ectotherms may be better
predicted by a curve that approaches the zero-line asymptotically
than by a single straight line (Peterson et al., 1983).
    Population growth comparisons. Due to an inability to measure
juvenile growth in the Spring, Gasconade and Big Piney Rivers,
the comparisons in Fig. 4 should be viewed with caution. The
lack of juvenile growth in the Gasconade is probably responsible
for the difference in slopes between the Gasconade and North Fork
hellbenders (Table 16). Because of the similar genetic frequencies
found by Merkle et al. (1977), it is reasonable that the amount
of growth at a specific length differs (elevations different)
between most populations (Table 16) because of differing environmental
conditions in different rivers. Selection can not be dismissed,
however, nor can founder events and drift, especially as gene
flow is probably impossible betwen populations, other than by
human intervention, except between the Gasconade and Big Piney
hellbenders.
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If only adult growth is considered, based on extensive observations of individuals from all six rivers, the comparison of growth in Fig. 4 seems reasonable. The Niangua and Spring River adults attain the largestsize followed by the Gasconade hellbenders. The North Fork and Big Piney adults seem similar in size, and the Eleven Point hellbenders are much smaller than the others. This is indicated in the average weight of hellbenders in Table 8. Peterson et al. (1983) determined an average weight of 314 g for North Fork hellbenders based on over 700 individuals. Nickerson and Mâys (1973a) calculated an average weight of 671 g for Spring River hellbenders $(\mathrm{n}=33)$, 193 g for Eleven Point salamanders ( $\mathrm{n}=31$ ), 365 g for North Fork hellbenders $(\mathrm{n}=435)$, and 777 g for Niangua hellbenders $(\mathrm{n}=30)$. All but the North Fork animals were measured by the latter authors in the laboratory after regurgitation may have occurred. Juveniles compose a greater proportion of the population in the North Fork (unpublished data) and in the Eleven Point River and thus reduces the estimate of average weight in comparison to the other populations.

Environmental Factors Influencing Growth
Temperature. When food was not limiting, higher growth rates of Bufo boreas have been correlated with an optimum temperature because of the influence of temperature on ingestion rates and conversion efficiencies (Lilywhite et al., 1973). Highton (1962) indicated that the length of the growing season may be involved in the slower growth of slimy salamanders, Plethodon glutinosus, in Maryland as compared to a population in Florida. Temperatures are different among hellbender streams in the Ozarks both summer
and winter, exhibit different ranges of temperatures, and the optimum period of growth may differ slightly between north-flowing rivers and the more southern south-flowing streams. Lowry (1953) found that, in general, smallmouth bass growth was greater in north-flowing Ozark rivers than in south-flowing rivers. Yet no correspondence is noticeable between: growth patterns of hellbenders (Fig. 4) and temperature (Figs. 6 and 7). However, any interaction between growth and temperature may be too subtle to detect from: the data gathered in this study.

Food availability. Differential availability and quality of food has been suggested as a cause of differential rates of growth in different populations of turtles (Parmenter, 1980; Cagle, 1946) and lizards (Dunham, 1978; Schoener and Schoener, 1978; Andrews, 1976). A lower density population of spadefoot toads, Scaphiopus holbrooki, exhibited greater growth than a more crowded population (Pearson, 1955). Although there is no indication of a linear trend between weight of crayfish caught per hour and hellbender biomass, a weak correlation between hellbender biomass and the number of crayfish caught per hour exists among collection sites. Also a significant correlation was found between hellbender production estimates and the number of crayfish caught (Fig. 11). Further support of a relationship between hellbender production and crayfish production is given by the significant correlations produced when half Topping's (1981) average estimate of Niangua hellbender biomass (and production) is paired with the Niangua crayfish catch and included in the analyses. Even when divided in half, the Niangua average standing crop estimate is larger

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than for any other site. Yet the catch per effort of hellbenders
in the area where crayfish were caught in the Niangua is higher
than for any single catch in any other river. This suggests a
high biomass, particularly as hellbender growth may be greatest
in the Niangua River (Fig. 4).
    No cause and effect is assumed in correlation analysis however.
The relationship may be due to a third factor, such as the density
of cover items, because both animals hide beneath rocks during
the day. Further, with the difficulty in accurately estimating
population size, the correlation may be spurious. Also the density
of crayfish may vary seasonally. No attempt was made to measure
the availability of other food items. At most the analyses indicate
that crayfish abundance could be an important environmental factor
influencing the growth of hellbenders in different sites and suggest:
that it would have been beneficial to have obtained sufficient
data for growth comparisons within rivers. Further as suggested
by Lowry (1953) for smallmouth bass, Missouri fish growth often
varies more between portions of the same stream (headwaters compared
to lower portions) than between similar habitat types of different
streams (Purkett, 1958).
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                                    Age
    Length-age relationships. Although metamorphosis has been
reported to occur at }125\textrm{mm}\textrm{TL}\mathrm{ and }18\mathrm{ months of age, length may
vary at transformation just as at sexual maturity between populations.
This could lead to an erroneous length-age curve for Eleven Point
hellbenders. Bishop (1943) reported that hellbenders transform
at }100\mathrm{ to }130\textrm{mm}\mathrm{ TL and 18 months of age. However, even if 100
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$m m$ and 18 months is used as the reference point in the integrated equation relating age and $T L$, the constant of integration changes only from 119 to 126 . This produces a length-age curve only about seven months to the right of that in Fig. 8 for Eleven Point hellbenders.

Similarly, whether the actual size at sexual maturity is 360 or 380 mm TL changes the length-age curve for the other three populations by a maximum of $\pm 9$ months. Yet the method of aging is crude, particularly for older animals (cf. Westrhiem and Ricker, 1978) because of the variability of individual growth. Therefore no age class used in the fecundity and life tables is less than a five-year interval, and a very large interval is used for the oldest hellbenders.

The evidence is that maturity is more a function of age than size in hellbenders. Eleven Point females mature when about 7 to 8 years old at about 300 mm TL . North Fork and Niangua females also mature when about 7 to 8 years old, but at a length of about 330 to 380 mm for the former and 380 mm for the latter (Peterson et al., 1983; Taber et al., 1975). According to Wilbur and Collins (1973), amphibian metamorphosis, another life-history transition, is neither size nor age dependent alone, but is predicted best by the differentiation rate. Smith-Gill and Berven (1979) stated that $95 \%$ of the variance in the length of amphibian larval periods is accounted for by differentiation rates; only $50 \%$ of the variance is explained by growth rates. Nevertheless, Berven and Gill (1983) showed that Rana sylvatica tadpoles from Manitoba metamorphosed at variable sizes but at a constant age. In contrast, a Maryland
population responded with a constant size but variable age at metamorphosis. Policansky (1983) suggested that when platyfish, Xiphophorus maculatus, grow fast there is a minimum age that must be attained before sexual maturity occurs; however, when growth is slow, sexual maturity does not occur until a minimum weight is reached. Policansky (1983) postulated that if reproduction when small costs a lot in terms of survival and later reproduction and if reproductive success increases quickly with increased size, then size should be the trigger, at least in conditions that favor rapid growth. Alternatively, if reproductive success does not increase much with increased size and if mortality of reproductives of all ages is high, then the transition should be determined by age. However, neither of these alternatives seems to fit the life-history traits of Cryptobranchus well. Policansky (1983) did suggest that if the environment is stable in time and space, then there is no difference between age- and size-triggered changes. This is because in such an environment there will be little variation in growth rate for a given genotype, and age and size will be closely correlated. According to Williams et al. (1981), the hellbender is confined to a narrow niche by several adaptations of structure, behavior, and physiology brought about by its evolution in a relatively stable environment. However, Ozark streams differ environmentally at least somewhat with regard to temperature and possibly food availability. Further, hellbender growth is different despite little genetic differentiation among populations. Of course, no two habitats are exactly alike. Even different sites in the same river differ at least slightly in environmental factors and

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will differ in time.
If a life history transition is determined by the attainment of a certain age, conditions unfavorable to growth will result in a population of smaller individuals (Policansky, 1983), and individuals will have much lower fecundity if fecundity scales with size (Stearns, 1983a). Stearns (1983a) suggested that organisms should adopt an intermediate course between age-specific or size-specific transitions. He predicted qualitatively how age and length at maturity should change (plastic trajectories) when organisms from different populations subject to differing adult mortality are raised under conditions of food or temperature stress. If organisms have encountered high prematuration rates of mortality, then they should respond to stress by maturing at much smaller sizes with a small delay in maturity. Alternatively, if they have encountered relatively low prematuration mortality rates, then they should respond to stress with a relatively small reduction in size but a larger increase in age at maturity. Stearns' (1983a) trajectory for populations with a high prematuration mortality (m3, Fig. 3) is only slightly different than that of the fixed-age trajectory of maturation under "good," "normal," and "poor" environmental conditions. The growth models for the North Fork, Niangua, and Eleven Point populations may be too crude to detect such small differences in age at maturity and thus the populations could show such adaptations. Possibly the Niangua hellbenders with good environmental growth conditions are larger at maturity but slightly younger than the Eleven Point hellbenders experiencing poorer conditions for growth. The North Fork hell-
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benders (and presumably the other three populations) are in between
with regard to age and length at maturity and environmental conditions.
However, the concept is clouded somewhat because Stearns (1983a)
referred to prematuration mortality rates at one point, but to
adult mortality rates at another, in reference to his plastic
trajectories. For hellbenders the two are very different. Considering
the homozygosity found in hellbenders by Merkle et al. (1977),
such genetic changes may not be possible between populations.
Thus age may indeed be the important factor in triggering maturity
in all populations studied, while differing growth conditions,
such as food availability perhaps, determines the length at transi-
tions. Plasticity has been demonstrated in most life-history traits
(Caswell, 1983). Smith-Gill (1983) suggested that plasticity,
rather than being a genetic adaptation in many cases in higher
organisms, may be a passive response to a variable environment
and thus not necessarily adaptive. In fact, such variability
may be nonadaptive.
Maximum age. Because the growth of old individuals is so slow, the growth models cannot predict the maximum age of hellbenders. Fifty or more years does not seem unreasonable. Taber et al. (1975) reported that hellbenders live \(30^{+}\)years. One hellbender was maintained in captivity for 29 years (Oliver, 1955), and Flower (1936) stated that six Andrias japonicus, the largest extant salamander, have attained ages between 55 and 60 years in European 200s. It has been speculated that Andrias may live more than a century (Federal Writers' Project, 1939). Even much smaller urodeles have lived long at least in captivity. For example,
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Triturus cristatus (Biegler, 1966) and Ambystoma maculatum (Oliver, 1955) have lived 25 years and Salamandra salamandra 43 years (Schmidtler and Schmidtler, 1969). Hairston (1983) suggested that plethodon jordani has a life span of as much as 30 years under natural conditions. Further, a hellbender marked by Taber et al. (1975) and suspected to be $30^{+}$years old in 1972, was recaptured eight years later in apparent good health (Peterson et al., 1985). Robert Wilkinson (personal communication) believes that some hellbenders may attain an age of 75 years.
Age-specific Fecundity and Survivorship
Fecundity tables, Similar to the method used by Organ (1961)
to determine the number of individuals in the zero-age class for
Desmognathus, eggs were counted and a l:1 sex ratio at fertiliza-
tion was assumed. Also, because fertilization is external in
hellbenders, it was necessary to assume that all eggs were ferti-
lized to construct the fecundity tables (Table 20 ). Smith (1912)
found few hellbender eggs laid that were not fertilized, and
Caughley (1977) stated that vertebrates have a sex ratio so close
to 1:1 at birth that the slight preponderance of males can be
ignored. However, as mentioned previously, Topping and Ingersol
(1981) speculated that as much as $50 \%$ of the eggs produced by
Niangua females may be resorbed.
Because the greatest number of eggs are produced by the
youngest adulti (7 to l6 years of age), particularly in the Eleven
Point and Big Piney Rivers, it is this segment that would be most
advantageously manipulated to control populations. This is
especially true because little can be done to improve larval and
juvenile survivorship without raising them away from the river. As long as the young adults survive, the populations should be stable. The removal of old adults or small numbers of the young adults should not affect the population adversely. This is especially true if, as suggested by Taber et al. (1975) and Nickerson and Mays (1973a), cannibalism and egg-eating are a factor in population control. Because of the number of eggs eaten and because larvae and small juveniles are probably eaten when opportunity permits, it may be that this is the major population control. Although many animals in the rivers would eat the eggs, larvae, and small juveniles, adult hellbenders probably have greater opportunity because of occupying similar cover. Smith (1907) even reported that a captured $12-\mathrm{cm}$ TL larva regurgitated a $6-\mathrm{cm}$ larva. As suggested by Williams et al. (1981), protection of the habitat and prevention of commercial exploitation should allow for viable populations in the Ozarks.

Life tables. The age distributions of females were calculated by sampling within $\pm 2$ months of the birth pulse assuming a stable age distribution and a zero rate of increase. The frequency of the zero-age class is based on fecundity rates. Caughley (1977) lists three primary pitfalls for this type of life table based on a standing age distribution: (1) rate of increase must be known with reasonable accuracy, and this rate must have remained relatively constant for two or three generations; (2) although no absolute judgement on minimum sample size is possible because it depends on the accuracy required, a table based on less than 150 determinations of age is unlikely to be accurate enough for
any purpose when age distribution methods are employed; and (3) some indices of age have wide variances, and although they may be useful for dividing a sample into broad categories, they cannot be used in life table analyses. Information supporting a zero rate of increase has been discussed previously. The number of adults aged ranged from 66 to 99. However, Caughley (1977) presented a life table (Table 8.6) where 120 thar, Hemitragus jemlahicus, were aged. Only the zero-age class estimated from fecundity rates resulted in a sample of more than 150 age determinations. If this is acceptable then certainly the number of age determinations are sufficient in Table 21. It may seem unreasonable that hellbenders can be assigned to the age classes in the life tables with any certainty because of the inexactness of the aging method, particularly in the Spring River sample. But even if hellbenders are grouped into three broad categories, (young adults 7 to 16 years old, adults 17 to $27^{+}$years old, and old adults $27^{+}$years old), which seems reasonable, the pattern of survivorship for each population would be the same and would still forecast very high early mortality but increasing survivorship after sexual maturity.

All four populations exhibit the Type III survivorship curve, which represents the most extreme type of selection according to Pearl and Minor (1935). This fits the pattern of greater early mortality in a more aquatic urodele. Even the most aquatic species of Desmognathus studied by Organ (1961) exhibited a constant rate of survivorship with age or Type II survivorship curve.

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The two populations where growth is less, the Big Piney and the Eleven Point, exhibit slightly less juvenile mortality but slightly greater adult mortality than the other two populations in which growth is greater. Because there may be a bias toward older hellbenders in the Big Piney cohort due to heterogeneous catchability (Table 5), a survivorship curve even more like that of the Eleven Point females may be more realistic. Likewise, age-specific fecundity is less for the Eleven Point and Big Piney females than for the Spring and Gasconade females (Fig. 10). The greater fecundity might be an adaptation that has been selected because of the greater juvenile mortality of Spring and Gasconade River hellbenders. Yet the length-specific fecundity relationships (Fig. 2) strongly suggest that the pattern of fecundity in Fig. 10 is linked with growth. Even the small differences in survivorship that may exist between populations could be tied to growth. The larger Spring and Gasconade hellbenders would probably be less subject to predation than the smaller Big Piney and particularly the Eleven Point hellbenders as adults. Perhaps greater cannibalism and egg-eating because of larger size is a factor in the greater early mortality in the former two populations.

Natural selection has produced a species well adapted to large cool rocky streams having survived in such environments since the Pleistocene (Naylor, 1981). But little genetic variability remains (Merkle et al., 1977) The pattern of survivorship and fecundity differs significantly among hellbender populations. Although Stearns (1980) suggested that contrasts in life-history tactics are more evident at higher taxonomic levels and confirmed


#### Abstract

this for squamates (Stearns, 1984), Brown (1983) argued that intraspecific studies can provide greater understanding of the relative importance of the "proximal" causes of life-history variation. These include responses to variation in ecological properties, developmental plasticity, and genetic divergence resulting from selection or drift.

Growth appears to be the important factor causing differences in age-specific fecundity, and possibly survivorship, in different hellbender populations. Probably this is because of the plasticity of growth and the differences in environments in different rivers, such as food availability. According to Stearns (1984, 1983b), patterns of covariance of life-history traits in lizards, snakes, and mammals are strongly influenced by classwide correlations with a single trait, average adult female length; they are further affected by family and genus influences. Thus much of the tendency for traits to covary in the pattern originally thought to have been produced by $r$ - and K -selection can be explained by selection on size alone, followed by coadaptative shifts in life-history traits (Stearns, 1984). Characteristics that compose an organisms's life history are coadapted and modified by selection to solve a particular ecological problem (Stearns, 1976), implying that life-history characteristics have a genetic basis and have evolved under the influence of demographic selection (Berven and Gill, 1983). None the less, small differences in life history between populations should not be attributed to local evolution as a matter of course. In many cases differences may merely be a product


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of different environments influencing development passively and
may even be nonadaptive (Smith-Gill, 1983). Stearns (1977) pointed
out that a large number of studies only document variation in
life-history patterns between or within species but then comment
on its adaptive significance. Ignoring the variation induced
by the environment may lead to false conclusions (Berven, 1982).
This is true in reverse also; despite the evidence of similar
genetic frequencies in hellbenders, long-term transplant studies
must be conducted before these speculations on plasticity can
be examined fully.
    Hellbender density', growth, and age-specific fecundity and
survivorship are about as variable within populations of a subspecies
as between subspecies in the Ozarks. Among demographic traits
only length-specific fecundity is well segregated between subspecies:
a greater length-specific production of ova by all C. a. bishopi
populations (Table 13 and Fig. 2). However, even the effect of
this characteristic on life history may be masked by differences
in growth. It seems unlikely that the difference in demography
between the two subspecies in the Ozarks is any more profound
than what might be detected between isolated drainages in other
parts of the range of Cryptobranchus. Therefore demographic com-
parisons do not support the existence of subspecies.
                                    Female Cohort Biomass and Production
    The zero-age class in Table 22 represents half the expected
number of eggs produced annually by females in each population.
The high early mortality maintains a relatively low cohort biomass
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despite the large number of eggs. It is interesting that the total weight at death is supposedly greater for the Big Piney females than for the Gasconade females. This is due to the greater young mortality in the Gasconade, despite greater fecundity and adult survivorship. Because of the lack of data on juvenile survivorship, the calculated total biomass furnished to the predators and decomposers by a female cohort is possibly a minimum estimate. However, most of the mortality probably does occur in the egg or embryo stage, and not all eggs may be laid. Further, it is possible that a sizeable portion of the biomass is consumed by the hellbenders because of oophagy and cannibalism. The biomass of a male cohort would likely be less because of greater adult mortality (Peterson et al., 1983; Taber et al., 1975) and smaller size than the females.

Topping (personal communication) estimated annual hellbender production by a different measure of instantaneous growth. His estimates for 10 sites in the Niangua River varied from 60.7 to 714 (the next highest was 303) kg / ha with an average of 179 $\mathrm{kg} /$ ha. $\mathrm{P} / \mathrm{B}$ values varied from 0.16 to 0.22 . Topping's estimates for the most part are in accord with those made by the summationremoval method in Table 23 if male production is nearly comparable to that of females (and taking into consideration that only diurnal habitat was measured in the Niangua River). This supports the assumption of a stable population made in the construction of the life tables because no such assumption was necessary in Topping's estimates of production. Yet his $P / B$ values are very much in

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agreement with the ratios obtained by the instantaneous growth
method in Table 23, by which the estimates of yearly production
are much less. It may be that those estimates (P') are reduced
because of the overestimation of maximum age used as a divisor
in the calculations, or because the estimates of density and thus
biomass are reduced because of nonrandom sampling. However, such
P/B values might be expected because of the longevity of hellbenders.
According to Waters (1977), P/B ratios are reasonably constant
for given groups of organisms: modes of published annual P/B
values are 5.0 for univoltine zoobenthos, 1.2 for stream salmonid
fishes (effective life span three years), and 0.6 for warm-water
fishes (life span 10 years or more).
    If the estimates of production by the summation-removal method
are reliable, then hellbenders are likely an important part of
the stream trophic dynamics, at least in the areas where suitable
habitat exists. Waters (1983) estimated the annual production
of trout in a Minnesota stream for 15 years as brown trout, Salmo
trutta, replaced brook trout, Salvelinus fontinalis. His estimates
of production over that period ranged from 25 to 167, 3 to 45,
and 9 to 132 kg / ha / yr for brook, rainbow (Salmo gairdneri),
and brown trout, respectively. Covington et al. (1983) estimated
annual production of smallmouth bass, Micropterus dolomieui, and
rock bass Ambloplites rupestris, at two sites in the Jacks Fork
River and two sites in the Current River, both in the Ozarks.
Estimates varied from 2 to 16 kg / ha / yr for smallmouth bass
and from 6 to 22 kg/ ha / yr for rock bass. Production of small-
mouth bass was greater in the Jacks Fork, but rock bass production
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was greater in the upstream sites on both rivers. At the same
sites sampled by Covington et al. (1983), See (1982) estimated
production of black redhorse, Moxostoma duquesnei, age three and
older: 19.34 to 54.69 kg / ha / yr with the average production
greater in the Current River. Along with other production estimates
cited previously, these estimates suggest that hellbender produc-
tion may be of the same order of magnitude of many fish species
in the rivers, even if the instantaneous growth estimates of produc-
tion (P') are used for comparison instead of the higher production
estimates made by the summation-removal method (Table 23).
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            SUMMARY
    Estimates of hellbender densities ranged from 1.0 to 6.3
per 100 m}\mp@subsup{\textrm{m}}{}{2}\mathrm{ for four different populations and seven collection
sites (Table 24). Biomasses based on these estimates (and average
weights) varied from 59 to 396 kg / ha wet weight. These may
be minimum estimates because of bias produced by unequal catcha-
bility. Site areas were measured from shoreline to shoreline
and not only where diurnal habitat existed.
    Fecundity was positively correlated with female body length
in all populations, and C. a. bishopi populations exhibited greater
length-specific fecundity than C. a. alleganiensis populations.
Numbers of eggs removed from females that were to be laid during
the current breeding season varied from 296 to 908 ( }\overline{X}=576\mathrm{ ),
215 to 452 ( }\overline{X}=289),159 to 687(\overline{X}=374), and 95 to 481 ( \vec{X}=320
for Spring, Eleven Point, Gasconade, and Big Piney females, respec-
tively.
    Crayfish was the most common food item in stomachs in all
populations, but large numbers of hellbender eggs were probably
eaten during spawning.
    No significant linear relationship existed between wet weight
of crayfish caught per hour and hellbender production in different
collectionsites, but asignificant positive correlation was found
between the number of crayfish caught per hour and hellbender
production.
Growth decreases linearly with increased size of hellbenders. Growth of hellbenders is greatest in the Spring River followed
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Table 24. Summary of demographic parameters.

| Population | Average weight (g) | $\begin{gathered} \text { Density } \\ \left(\mathrm{no} . / 100 \mathrm{~m}^{2}\right) \end{gathered}$ | Biomass <br> (kg/ha) | Production by instantaneous growth method (kg/ha/yr) | $\begin{array}{r} \begin{array}{r} \text { Aver } \\ \text { crayfish } \end{array} \\ \hline \text { (no./hr) } \end{array}$ | $\begin{aligned} & \text { ge } \\ & \text { catch } \\ & (\mathrm{g} / \mathrm{hr}) \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\begin{gathered} \text { c. an. } \\ \text { bishopi } \end{gathered}$ |  |  |  |  |  |  |
| Spring 939617781792 |  |  |  |  |  |  |
| site 2 | 769.2 | 1.0 | 74 | 14.0 | 127 | 888 |
| Eleven Point |  |  |  |  |  |  |
| site 2 | 346.0 | 6.1 | 209 | 36.7 | 92 | 729 |
| $\begin{gathered} \text { C. a. } \\ \text { alleganiensis } \end{gathered}$ | \% |  |  |  |  |  |
| Gasconade site 1 | 597.6 | 1.0 | 59 | 11.2 | 50 | 349 |
| Big Piney 20.30 .388 |  |  |  |  |  |  |
| site 1 | 486.7 | 2.3 | 112 | 20.3 |  |  |
| site 2 | 531.5 | 6.3 | 334 | 60.4 | 153 | 923 |

Table 24 cont. Summary of demographic parameters.

| Population | Population composition average \% |  |  | ```Average no. female zygotes produced per female by females aged in years \[ (7-16) \quad(17-26) \quad\left(27^{+}\right) \]``` |  |  | Type survivorship curve | Mean generation <br> time (yrs) | Production of females by summationremoval method (kg/ha/yr) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | juveniles | adult <br> males | $\begin{aligned} & \text { adult } \\ & \text { females } \end{aligned}$ |  |  |  |  |  |  |
| $\begin{gathered} \text { c. a. } \\ \text { bishopi } \end{gathered}$ |  |  |  |  |  |  |  |  |  |
| Spring | 0.5 | 50.0 | 49.5 | 236 | 323 | 362 | III | 23.3 | 96.5 |
| Eleven Point | 24.8 | 25.0 | 50.2 | 125 | 174 | 186 | III | 14.2 | 92.4 |
| $\begin{aligned} & \text { C. a. } \\ & \text { alleganiensis } \end{aligned}$ |  |  |  |  |  |  |  |  |  |
| Gasconade | 0.5 | 36.0 | 63.5 | 147 | 216 | 243 | III | 23.4 | 22.6 |
| Big Piney | 5.3 | 39.7 | 54.9 | 142 | 201 | 218 | III | 18.6 | 73.1 |

```
by the Gasconade and Big Piney hellbenders; growth is least in
the Eleven Point River. Water quality reports indicate that the
Spring River may be nutrient enriched compared to the Eleven
Point River.
    Age-specific fecundity of hellbenders in different populations
is also greatest in the Spring River followed by the Gasconade,
Big Piney, and Eleven Point populations. It is believed that
differences in growth, due mostly to environmental differences
among rivers, highly influences fecundity. Assuming a stationary
age distribution, a Type III survivorship curve is exhibited by
all populations of the strictly aquatic hellbender. There is
little demographic data to support the existence of subspecies.
    Estimates of annual production indicate that hellbenders
may at least be important in the trophic relationships of the
streams in the circumscribed area of suitable habitat.
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[^0]:    ${ }^{\text {a }}$ Leslie suggests that number known alive should exceed 20

[^1]:    *, **, *** $\mathrm{P}<0.1,0.05,0.01$ respectively

[^2]:    *, **, *** $\mathrm{P}<0.1,0.05,0.01$ respectively

[^3]:    *, **, *** $\mathrm{P}<0.1,0.05,0.01$ respectively

[^4]:    $\mathrm{a}_{\text {Estimate }}$ calculated by method described by Caughley (1977).

[^5]:    $a_{\text {Estimate calculated by method described by Caughley (1977). }}$

[^6]:    *, **, *** $\mathrm{P}<0.1,0.05,0.01$ respectively

[^7]:    *, **, *** $\mathrm{P}<0.05,0.01,0.001$ respectively

[^8]:    *, **, *** $\mathrm{P}<0.05,0.01,0.001$ respectively

[^9]:    *, **, *** $\mathrm{P}<0.05,0.01,0.001$ respectively

[^10]:    $a_{\text {Assumes }}$ a $1: 1$ zygote sex ratio and that all eggs are laid and fertilized

[^11]:    ${ }^{\text {a }}$ Assumes a $1: 1$ zygote sex ratio and that all eggs are laid and fertilized.

[^12]:    $a_{\text {From }}$ fecundity table

[^13]:    ${ }^{a}$ From fecundity table

[^14]:    ${ }^{\text {Approximate }}$ weight of an egg before laying
    $I x=W x-1+((N x / N x-1) \cdot G x) ; T x=D x \cdot I x$ from Spight (1967)

[^15]:    $a_{\text {Approximate }}$ weight of an egg before laying
    $I x=W x-1+((N x / N x-1) \cdot G x) ; T x=D x \cdot I x$ from Spight (1967)

[^16]:    ${ }^{\text {a }}$ Approximate weight of an egg before laying
    $I x=W x-1((N x / N x-1) \cdot G x) ; T x=D x \cdot I x$ from Spight (1967)

